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Investigation of Influential Variables to Predict Passing Rate at Short Passing Zones on Two-Lane Rural Highways

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Investigation of Influential Variables to Predict Passing Rate at Short Passing Zones on Two-Lane Rural Highways / Karimi, Arastoo; Mirza Boroujerdian, Amin; Bassani, Marco. - In: JOURNAL OF TRANSPORTATION ENGINEERING. - ISSN 2473-2907. - ELETTRONICO. - 146:10(2020), p. 04020117. [10.1061/JTEPBS.0000440]

Availability: This version is available at: 11583/2842736 since: 2020-08-18T10:41:32Z

Publisher: American Society of Civil Engineers

Published DOI:10.1061/JTEPBS.0000440

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1	Investigation of Influential Variables to Predict Passing Rate at Short
2	Passing Zones on Two-Lane Rural Highways
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17 Abstract

18 The passing zone (PZ) is that part of two-lane highways in which drivers can safely overtake 19 slower vehicles. Several studies have presented passing rate models in the PZ. However, there is 20 no model to predict the passing rate in PZs shorter than 350 m. Furthermore, the effect of variables 21 such as lane width, and the proportion of motorcycles on the passing rate were not investigated in 22 previous works. This study assessed the effects of these variables on the passing rate and present 23 a prediction passing model for short PZs. Data were collected from seven PZs using a drone on 24 three different two-lane rural highway segments in Iran. The results showed that the passing rate 25 depends on the lane width, absolute vertical grade, the flow rate in both directions, directional 26 split, the proportion of heavy vehicles in the subject direction, and the proportion of motorcycles 27 in the subject direction. Short PZ length values did not have a significant effect on the passing rate. 28 The passing capacity occurred at a flow rate of 680 veh/h in both directions irrespective of the 29 directional split.

30

Keywords: passing rate, short passing zone, two-lane rural highway, lane width, the proportion of
motorcycles.

33 **1. Introduction**

34 On two-lane highways, faster vehicles are forced to follow slow vehicles until they reach a stretch 35 of road which provides a passing opportunity, which depends on a gap in oncoming traffic and 36 sight distance. The passing zone (PZ) provides a sight distance value which is sufficient for passing 37 maneuvers. Where many passing maneuvers are possible, the traffic operation performance will 38 improve. Passing demands and passing capacity (maximum passing rate) have a considerable 39 impact on operation and driver perception of service. The passing rate indicates the number of 40 passing maneuvers conducted in an hour by vehicles travelling in the same direction on a two-lane 41 road.

42 The Highway Capacity Manual (HCM) (TRB, 2010) used the percent time spent following 43 (PTSF) as a surrogate measure for passing demand to measure the operating performance of 44 two-lane rural highways. However, HCM does not use the passing rate in the operation analysis 45 because of the lack of comprehensive studies in this area. Passing rate modelling requires accurate 46 observational field data: each individual vehicle needs to be constantly monitored while engaged 47 in passing maneuvers. Hence, field data are difficult to collect due to the considerable length of 48 PZs. For example, Mwesige, et al. (2016) used pneumatic tubes and camcorders mounted on 49 tripods placed by the roadside along the PZ. Moreno, et al. (2013) used a mobile traffic laboratory 50 which was equipped with six digital video cameras installed on an elevated platform.

Harwood, et al. (2008) presented two categories for the PZs: short PZs (PZ lengths shorter than 240 m) and long PZs (PZ lengths equal to 300 m or more). Using traffic simulation analysis, they showed that very few passing maneuvers occurred in short PZs. Their field observations also showed that only 0.4% of all vehicles made passing maneuvers in short PZs while 92% of passing maneuvers ended in no-passing zones (NPZ). Hence, their conclusion that short PZs contribute 56 little to operational efficiency. However, they did not develop the passing rate prediction model 57 for PZs. While a few field studies investigated the passing rate models (Mwesige, et al., 2016, 58 Moreno, et al., 2013, Hegeman, 2008), none of them considered short PZs in their analysis. 59 Furthermore, the effects of two explanatory variables - lane width and proportion of motorcycles 60 on the passing rate - were not investigated in previous works.

61 To address this gap in knowledge, this study has investigated the effects of geometry and 62 traffic-related variables, including the lane width and proportion of motorcycles, on the passing 63 rate in PZs shorter than 350 m. Videos of vehicles involved in passing maneuvers were collected 64 along short PZs using a drone. Video analyses were carried out to obtain independent variables 65 that were then used to develop a first passing rate model for those maneuvers ending in the PZ 66 (i.e., PZ-PZ and NPZ-PZ), and a second passing rate model for those which both start and end in 67 the PZ (i.e., PZ-PZ). The two models are appropriate for countries where it is legally accepted that 68 passing maneuvers may start in the NPZ and end in the PZ (NPZ-PZ), and to countries where they 69 must start and end in the PZ (PZ-PZ) (Mwesige et al., 2016).

The manuscript is organized as follows: Section 2 presents a literature review on passing frequency along PZs; Section 3 explains how to model the passing rate statistically and also introduces the study sites and data collection method; Section 4 provides the results of the study, while the discussion of the results is presented in Section 5; Section 6 provides the main findings and addresses implications, recommendations, and future work needs.

75

76 2. Background

Wardrop (1952) offered a theoretical model to estimate passing demand from the speeddistribution while assuming an ideal situation with no opposing traffic and no passing sight

79 limitations. He used eq.(1) to calculate the number of passing maneuvers per kilometer per hour 80 (P_n) :

81
$$P_n = 5.6 \left(\frac{V_s^2 \cdot \sigma_s}{\overline{\mu}_s^2} \right)$$
(1)

82 where V_s is traffic flow in the subject direction [veh/h], $\overline{\mu}_s$ is the average of space mean speeds 83 [km/h] and σ_s is the standard deviation of space mean speed [km/h].

Daganzo (1975) developed a theoretical negative exponential passing rate model based on traffic flow in both directions to consider the reduction in passing opportunities due to oncoming traffic. He assumed that traffic flow and speed values were the same in both directions.

McLean (1989) proposed a formula for passing opportunities which used both oncoming
traffic and sight distance limitations as follow:

$$P(o.t) = P(g > 30) \times P(road)$$
(2)

90 where P(o.t) is the probability of an overtaking opportunity, P(g>30) is the proportion of time in 91 which there is a gap larger than the critical gap (30 s), and *P(road)* is the proportion of road length 92 which is suitable for passing. In McLean's method, a passing maneuver is possible when the 93 opposite gap is larger than 30 s (critical gap). He calculated P(g>30) as a function of the proportion 94 of following vehicles and mean free headway. McLean (1989) determined P(road) as a function 95 of percentage of PZs length along the road segment, frequency of NPZs, and speed of traffic flow. 96 Dommerholt and Botma (1988) attempted to capture both passing opportunities and 97 passing demand for the development of a passing rate model. They developed the theoretical 98 passing rate as a function of the standard deviation of the speed and the density of vehicles. To 99 calculate the *expected number of overtaking maneuvers*, they reduced the theoretical passing rate 100 using the probability that the passing maneuver was possible. To verify their proposed model,

Dommerholt and Botma (1988) investigated six locations. Real passing rates were lower than the values obtained from the Dommerholt and Botma equation. To reduce the theoretical passing rate, they took the following two factors into account: (1) that the passing maneuver does not occur if the speed difference between the lead and following vehicles is below a threshold value, and (2) that the passing rate reduces linearly with increasing mean platoon length for traffic flows above 400 veh/h.

107 More recently, Tuovinen and Enberg (2006) developed separate linear regression models 108 for individual road segments in order to estimate passing rates. One-way traffic flows were used 109 as the only predictor variables in the quadratic form. Hegeman (2008) studied passing maneuvers 110 along two road segments, one of which had a passing prohibition. This prohibition was applied 111 according to the Dutch Sustainable Safety Program (Wegman and Aarts, 2006), which suggested 112 the prohibition of passing maneuvers as a means to improve safety. Hegeman (2008) developed a 113 multivariate linear regression model to estimate passing rates based on subject and oncoming flow 114 rates:

115

$$OF = 1.6 \times 10^{-11} \times V_s^{1.5} (1700 - V_o)^{2.5}$$
(3)

where OF is the passing rate per kilometer per hour, Vs is traffic flow in the analysis direction [veh/h], V_O is traffic flow in the opposite direction [veh/h].

Moreno, et al. (2013) developed a Poisson regression model such as that in Equation (4) to
predict the passing frequency in the subject direction for a 15-min period (PF):

120
$$PF = \exp\left(\frac{-4.57904 - 0.00125 \cdot V^2 - 0.0000013 \cdot L_{PZ}^2}{+2.75645 \cdot D_S + 0.04093 \cdot V + 0.003455 \cdot L_{PZ}}\right)$$
(4)

where *V* is the two-way traffic volume for a 15-min period [veh/h], L_{PZ} is the length of PZ [m], and D_S is the directional split in the subject direction. They used two different road segments with different lengths of PZ in two directions and used 114 periods of 15 min per passing zone to
develop their model. The maximum passing rate occurred at a two-way traffic volume between
600 and 700 veh/h for all PZs.

126 Mwesige, et al. (2016) proposed a model to predict the passing rate per hour (*P*) in the 127 subject direction at PZs based on traffic and geometric parameters as follows:

128
$$P = \exp \begin{pmatrix} -5.089 + 1.123 \cdot L_{PZ} - 0.1948 \cdot L_{PZ}^2 + 0.09951 \cdot V_G \\ +0.01917 \cdot V - 0.00002401 \cdot V^2 + 0.008177 \cdot S_{85} \\ +0.0376 \cdot D_S + 0.05555 \cdot P_{HV} + 0.0008065 \cdot P_{HV}^2 \end{pmatrix}$$
(5)

where L_{PZ} is the length of PZ [km], V_G is the absolute vertical grade [%], V is the total traffic volume in the two travel directions [veh/h], S_{85} is the 85th percentile speed [km/h], D_S is the directional split in the subject direction, and finally P_{HV} is the proportion of heavy vehicles.

Mwesige, et al. (2016) used a total of 96 observations to estimate their models. They found that by increasing the percentage of heavy vehicles, the passing rate rose, then reached a peak of 35% before falling back. Based on their model, passing capacity occurred at the two-way volume of 400 veh/h. Other variables including the length of PZ, absolute vertical grade, and directional split in the subject direction were significant at the 95% confidence level. The 85th percentile speed was significant at the 90% confidence level. Table 1 presents a list of the above-mentioned studies and their corresponding contributions (e.g., explanatory variables) and shortcomings.

A number of field studies were conducted to find the influential variables on the passing rate in PZs. No study presented a passing rate prediction model for short PZs. Furthermore, the effects of both the proportion of motorcycles and lane width variables have yet to be investigated.

143 **3. Objectives, methodology, and data collection**

144 3.1 Objectives

This research aims to determine the effects of geometric (i.e., short PZ length, lane width) and traffic-related variables (i.e., the proportion of motorcycles) on the passing rate and the passing capacity. This research used field data collected with a drone to develop models for predicting the passing rate in the PZs. In particular, two models were developed: (i) Model 1 estimates the number of passing maneuvers in 15 minutes ending in PZ (PZ-PZ and NPZ-PZ) in the subject direction, while (ii) Model 2 estimates the number of passing maneuvers in 15 minutes which both start and end in the PZ in the subject direction (PZ-PZ).

152

153 **3.1. Variables**

154 **3.1.1 Dependent variable**

155 In most cases, passing maneuvers that start and end in the PZ (PZ-PZ maneuvers) have sufficient 156 sight distance (Khoury and Hobeika, 2007). Passing maneuvers that begin in the NPZ and end in 157 the PZ (NPZ-PZ maneuvers) are more likely to be initiated near the beginning of the PZ (Mwesige, 158 et al., 2016). Harwood, et al. (2008) categorized them as *jumping*. Although drivers do not have a 159 passing sight distance sufficient to complete the entire passing maneuver, it should be enough to 160 reach the abreast position, at which point they can abort the maneuver if necessary. Hence, drivers 161 have a sight distance which is sufficient for the risk evaluation of passing maneuvers ending in the 162 PZ (PZ-PZ and NPZ-PZ). However, many countries legally accept only those maneuvers that start 163 and end in the PZ (PZ-PZ), while others accept both cases (Mwesige, et al., 2016).

In light of the above, two dependent variables were considered: (i) the number of passing maneuvers ending in the PZ in the subject direction, and (ii) the number of passing maneuvers starting and ending in the PZ in the subject direction.

167

168 **3.1.2 Explanatory variables**

169 Explanatory variables were derived from both theoretical and field observation models described
170 in the Background section; however, some additional variables expected to be influential were also
171 examined.

Variables were categorized into two groups: (i) geometry-related variables like PZ length, upstream NPZ length, absolute vertical grade, and lane width, and (ii) traffic-related variables like traffic flow rate in both directions, traffic flow rate in the subject direction, traffic flow rate in the opposite direction, directional split of traffic volume, proportion of heavy vehicles, proportion of motorcycles, mean free-flow speed, standard deviation of free-flow speed, and 85th percentile of free-flow speed.

178

179 **3.2. Statistical analysis**

The number of passing maneuvers takes only non-negative integer values. Thus, in this case, the count data models provide an appropriate framework for estimating the number of passing maneuvers. Poisson regression is the most popular method for modeling count data. In the Poisson regression model, the probability of observation *i*-th having y_i completed passing maneuvers in 15-min, is given by:

185
$$P(y_i) = \frac{\exp(-\lambda_i)\lambda_i^{y_i}}{y_i!}$$
(6)

186 where λ_i is the Poisson parameter, which is equal to the number of expected passing maneuvers 187 for a *i*-th 15-min. One of the properties of a Poisson distribution is that the mean of the counting 188 process equals its variance. The Poisson regression model is estimated by specifying λ_i as a 189 function of explanatory variables, as shown in:

190
$$\lambda_i = E[y_i] = VAR[y_i] = \exp(\beta X_i)$$
(7)

191 where X_i is the vector of explanatory variables, and β is the vector of estimable parameters from 192 observed data. This model was estimated using the standard maximum likelihood method. The 193 likelihood function computed for all observations is:

194
$$L(\beta) = \prod_{i} \frac{\exp\left[-\exp\left(\beta X_{i}\right)\right] \left[\exp\left(\beta X_{i}\right)\right]^{y_{i}}}{y_{i}!}$$
(8)

The equality between mean and variance represents a limitation in the case of over-dispersed data (i.e., when the variance is greater than the mean). In many studies, the omitted variables are the main reason for overdispersion (Washington, et al., 2010). Overdispersion leads to inflation in estimated standard errors of estimated parameters, but it does not affect the magnitude of estimated parameters (Cameron and Trivedi, 2013). The Negative Binomial model is able to account for overdispersion, which is derived by rewriting Equation (7) into Equation (9) :

202
$$\lambda_i = \exp(\beta X_i + \varepsilon_i)$$
(9)

where $\exp(\varepsilon_i)$ is the disturbance term, which has a Gamma distribution with mean one and variance a. This term makes it possible to have a variance different from the mean, as shown in:

205
$$VAR[y_i] = E[y_i] + \alpha E[y_i]^2 = \lambda_i + \alpha \lambda_i^2$$
(10)

where α is the overdispersion parameter. When α is zero, the negative binomial and Poisson models are the same, therefore the choice between these two models depends on the value of α . The Negative Binomial probability distribution is:

209
$$P(y_i) = \frac{\Gamma((1/\alpha) + y_i)}{\Gamma(1/\alpha) y_i!} \left(\frac{1/\alpha}{(1/\alpha) + \lambda_i}\right)^{1/\alpha} \left(\frac{\lambda_i}{(1/\alpha) + \lambda_i}\right)^{y_i}$$
(11)

where $\Gamma(.)$ is the Gamma function. The parameters of the negative binomial model were able to be estimated using again the maximum likelihood method (Cameron and Trivedi, 2013).

In this study, the passing rate data were obtained from seven PZs along three different two-lane rural highways. Therefore, there may be a correlation among observations in each PZ because the data from each PZ may share unobserved effects. Random effects models should be considered to account for correlations of observation in each PZ. Random effects may be considered in count data models based on:

217
$$Ln(\lambda_{ij}) = \beta X_{ij} + \eta_j$$
(12)

where λ_{ij} is the expected number of completed passing maneuvers for the *i*-th 15-min belonging to *j*-th PZ, X_{ij} is a vector of explanatory variables, β is a vector of corresponding parameters, and η_j is a random effect for observations of the *j*-th PZ. Based on this specification, the random-effects Poisson model is:

222
$$P(y_{ij}|X_{ij},\eta_j) = \frac{\exp\left[-\exp\left(\beta X_{ij}\right)\exp\left(\eta_j\right)\right]\left[\exp\left(\beta X_{ij}\right)\right]}{y_{ij}!}$$
(13)

It is assumed that η_j is randomly distributed across groups (PZs) such that $\exp(\eta_j)$ has Gamma distribution with mean one and variance φ . For the random-effect Poisson model, the mean value is not equal to the variance, and the variance to mean ratio is $1 + \lambda_{ij}/(1/\varphi)$. If φ is zero, the random-effects and pooled (Poisson) models are not significantly different. The random-effects 227 Negative Binomial model can be derived using the same approach as above for the random-effects 228 Poisson model, which was described by Haussman, et al. (1984).

229

In this paper, the presence of overdispersion and correlation among observations was 230 examined, then the appropriate model was chosen to estimate the passing rate.

231

232 3.3. Study sites and data collection

233 Data were collected from seven PZs on three different two-lane rural highways in Iran 234 (Jiroft-Faryab, Jiroft-Baft, and Jiroft-Kerman highways). All passing zones were along straight 235 sections and almost constant vertical grades. Table 2 presents the geometric characteristics and 236 posted speed limits for PZs. PZ length varied from 164 to 345 meters. Lane width was between 3 237 and 3.75 m, and the absolute vertical grade was between 0.5 to 9.5%. PZ n.7 (Table 1) is the only 238 one with a shoulder, which is paved. Figure 1 shows diagrams of the horizontal and vertical 239 alignment around the investigated sections.

240 Data were collected using a Phantom 4 Pro drone (Figure 2) equipped with a 1-inch 241 20-megapixel sensor capable of shooting 4K/60fps video and a 3-axis gimbal to stabilize the 242 camera oscillation. During video recording, the minimum altitude of the drone was 150 m to avoid 243 any impact on driver behavior. Data were collected during 38 flights lasting 15-20 minutes each. 244 The weather conditions during data collecting were clear and good.

245 Using road markings at the beginning of the PZs (their lengths were measured in the field) 246 and the timestamps of vehicles determined by using the open-source video analysis software 247 Kinovea (Charmant, 2016), vehicle speeds and time headways were calculated. The vehicle type 248 was visually recorded. Free-flow speeds were determined for passenger cars that were found to 249 operate in free-flow conditions based on at least 6 s headway (Al-Kaisy and Karjala, 2008).

Geometry variables such as lane width, vertical grade, and PZ length were measured in the field.The hourly flow rate was calculated as the volume of vehicles in 15 minutes multiplied by four.

252

253 **4. Results**

254 Table 3 presents a summary of the observed variables at the PZs. The average number of passing 255 maneuvers ending in the PZ in the subject direction (N_{PZ}) was 5.0 passes per 15 minutes, reaching 256 a maximum value of 33 passes per 15 minutes. The average and maximum values of the number 257 of passing maneuvers starting and ending in PZ in the subject direction (N_{PPZ}) were 2.3 and 21, 258 respectively. The average of N_{PZ} was more than two times that of N_{PPZ} , which means that more 259 than half of the drivers ending their maneuver in PZ, started it from NPZ. The flow rate in both 260 directions had a range between 212 and 840 veh/h, with a directional split up to 80%, which provides a suitable combination for analysis. The directional split equals 30% means ($V_S = 0.3V$, 261 262 and $V_0 = 0.7V$). The proportion of motorcycles in the subject direction had a maximum value of 263 14%, which was a considerable one for rural highways. Mean free-flow speed for every 15 minutes 264 had an average of 85.3 km/h with an average standard deviation of 16.0 km/h. Based on the 265 explanatory variables presented in Table 3, the N_{PZ} and N_{PPZ} are estimated using count data models.

266 Table 4 presents a summary of the estimation of Models 1 (to estimate the N_{PZ}) and 2 (to estimate the N_{PPZ}). To find a suitable model for estimating the passing rate, overdispersion in the 267 data and correlations among observations in each PZs were analyzed. Table 4 shows the results of 268 269 the likelihood ratio test for the overdispersion parameter (α): it illustrates that α is statistically 270 highly insignificant for both Model 1 (p-value = 0.500) and Model 2 (p-value = 0.379). This means 271 that α is equal to zero, i.e., there is no difference between the Negative Binomial and Poisson 272 models. Table 4 also shows that the random-effects parameter (φ) is highly insignificant for both 273 Model 1 (p-value = 1.00) and Model 2 (p-value = 1.00). These results imply that the parameter of φ is statistically equal to zero, and there is no difference between the random-effects and pooled 274 275 (Poisson) models. Hence, the Poisson model was chosen to estimate the passing rate in this study. 276 The passing rate models were estimated using the STATA statistical software (StataCorp, 2017). 277 Different model forms using variables that were listed in Table 3 were estimated. In the following, 278 the results of model estimation for Models 1 and 2 are carefully described.

279

280 **4.1. Estimating** *N_{PZ}* (Model 1)

281 The overall significance of Model 1 was evaluated using the likelihood ratio test, which was significant at the 95% confidence level ($\chi^2 = 226.73$, *p*-value < 0.0001). Deviance and Pearson 282 283 tests were conducted to assess the goodness-of-fit of the model. The insignificant test statistics (p-value > 0.05) indicate that the model fits the data well. As additional descriptive measures of 284 goodness-of-fit, Cragg-Uhler R² (Cragg and Uhler, 1970) and McFadden's R² (McFadden, 1973) 285 were calculated (values equal to 0.950 and 0.418 respectively). These two statistics confirm that 286 287 the model fits the data well. Wald tests were carried out to identify significant variables which 288 were retained in the model.

289 Variables of the PZ length, upstream NPZ length, 85th Percentile free-flow speed, and 290 standard deviation of free-flow speed were statistically not significant at the confidence level of 291 90% and, therefore, removed from the model. However, the signs of the coefficients for these 292 variables were consistent with a priori expectation. The lane width was found to be highly 293 significant (*p*-value < 0.001). Its positive sign implies that an increase in lane width leads to an 294 increase in the passing rate. Table 5 shows that the average marginal effect of lane width on passing 295 rate at 15 min is equal to 7.47, which means that if the lane width increases by 0.2 m, N_{PZ} increases 296 by 6 passes per hour on average. The marginal effect of lane width increases by increasing its 297 value. The absolute vertical grade was statistically significant at the 95% confidence level (p-value 298 < 0.01). Based on its marginal effect reported in Table 5, each 1% increase in the absolute vertical 299 grade, the N_{PZ} increases by the hourly passing rate of 1.56 passes on average.

300 The flow rate in both directions (V) and its quadratic term had a significant effect at the 301 95% confidence level (*p*-values < 0.05). This variable has different effects based on its value. 302 Figure 3 illustrates the average marginal effects of flow rate in both directions on N_{PZ} for different 303 values of the flow rate. This figure shows that the positive marginal effect peaked at the flow rate 304 of 400 veh/h. The figure shows a sharp fall in the marginal effect of flow rate such that it was zero 305 at the flow rate of 680 veh/h. From this value, the effect of flow rate on N_{PZ} was negative, which 306 means that at an increased flow rate (greater than 680 veh/h), the passing rate decreased. The 307 directional split of traffic volume (D_S) was also a significant variable at the level of 0.001. With 308 an increase in Ds, the traffic in the subject direction increases and the traffic in the opposite 309 direction decreases, with the former resulting in an increase in the passing demand, and the latter 310 leading to an increase in passing opportunities. Another way to capture this effect in the models is 311 to use the two variables of the flow rate in the subject and flow rate in the opposite direction.

However, using flow rate and directional split variables presented better model results. The results show that with a 10% increase in directional split for subject direction, the N_{PZ} increased by 6.8 passes per hour on average.

The proportions of heavy vehicles (P_{HVS}) and motorcycles (P_{MCS}) to traffic volume in the subject direction were significant at the 95% and 90% confidence levels, respectively. The marginal effect of the P_{MCS} is almost twice that of the P_{HVS} . A five percent increase in P_{HVS} increased the N_{PZ} by 2 passes per hour on average. A five percent increase in P_{MCS} increases the passing rate by 4.2 passes per hour.

320

321 **4.2. Estimating** *N*_{*PPZ*} (Model 2)

Model 2 was significant at the 95% confidence level ($\chi^2 = 84.488$, p-value < 0.0001). The results of the Pearson goodness-of-fit test suggest the model performed well ($\chi^2 = 226.73$, *p*-value = 0.099). However, the Deviance statistic was significant ($\chi^2 = 90.559$, *p*-value = 0.042).

The length of the PZ had a statistically significant effect on N_{PPZ} at the 95% confidence level (*p*-value < 0.01). Table 6 indicates that the average marginal effects of the PZ length on N_{PPZ} were 0.018, which means if the L_{PZ} increased 100 m, the N_{PPZ} increased by 7.2 passes per hour. The variables L_W , V_G , and D_S were also statistically significant as they were for Model 1. Their average marginal effects are presented in Table 6. Unlike Model 1, the two variables of P_{HVS} and P_{MCS} were statistically insignificant at the 95% confidence level.

331

332 **5. Discussion**

The upstream NPZ length did not have a significant effect on the passing rate. Mwesige, et al.
(2016) reached a similar result on the effectiveness of the upstream NPZ length. They presented

four factors that could explain this result. The two variables of 85th percentile free-flow speed and standard deviation of free-flow speed, like the study of Mwesige, et al. (2016), were statistically not significant at the 95% confidence level.

338 According to previous works, PZ length is statistically significant (Mwesige, et al., 2016, 339 Moreno, et al., 2013). However, in this study, the PZ length had a strongly insignificant effect on 340 passing maneuvers ending in the PZ (N_{PZ}). The PZs that previous works studied were within the 341 range of 290-2990 m (Mwesige, et al., 2016) and 256-1270 m (Moreno, et al., 2013), respectively. 342 The PZs in this study ranged from 164 to 345 m, with most of them on the short side. The suspect 343 is that almost all drivers were familiar with the highways; hence they were aware of the short 344 length of the PZs and anticipated their maneuvers to complete them safely. The results presented 345 in Table 3 show that the average number of NPZ-PZ maneuvers was higher than that for the PZ-PZ ones. However, if only the N_{PPZ} were considered, the PZ length was significant. It may be 346 347 concluded that drivers (most of whom were familiar with the road) adjust the starting point of their 348 passing maneuver so as to complete it safely before the NPZ.

A sensitivity analysis of Model 1 explanatory variables was conducted to see how an explanatory variable affected the passing rate. To present the 15-min passing frequency as an hourly passing rate, N_{PZ} and N_{PPZ} are multiplied by four as equation (14) and (15) illustrate:

352
$$P_{PZ} = 4 \cdot N_{PZ} = 4 \cdot \exp\left(\frac{-8.4586 + 1.4989 \cdot L_W + 0.0779 \cdot V_G + 0.00852 \cdot V - 6.28 \times 10^{-6} \cdot V^2}{+0.03381 \cdot D_S + 0.02044 \cdot P_{HVS} + 0.04307 \cdot P_{MCS}}\right) (14)$$

353
$$P_{PPZ} = 4 \cdot N_{PPZ} = 4 \cdot \exp\left(\frac{-12.31318 + 0.00786 \cdot L_{PZ} + 1.71082 \cdot L_{W} + 0.06994 \cdot V_{G}}{+0.01218 \cdot V - 10.2 \times 10^{-6} \cdot V^{2} + 0.02636 \cdot D_{S}}\right)$$
(15)

Figure 4 produces for $V_G = 3.5\%$, $P_{HVS} = 8.56\%$, $D_S = 50$, $L_W = 3.45$, and $P_{MCS} = 0\%$ in equation (14), except the variables that change in each figure. 356 Figure 4a provides information on the change in the passing rate of maneuvers ending in 357 the PZ (P_{PZ}) for different values of the lane width at four different levels of the traffic flow rate 358 (200, 400, 600, and 800 veh/h). The results show that by increasing the lane width, the passing rate 359 increased. The marginal effect of lane width increased as the base value increased. For example, 360 adding 0.1 m to a lane with 3.6 m width (base value) has more effect on the passing rate than 361 adding 0.1 m to a lane with 3 m width. A wider lane width results in drivers taking more risks 362 because it provides more room for passing and helps to prevent opposing vehicles from colliding. 363 Furthermore, the marginal effect of lane width with respect to the flow rate is maximized at the 364 traffic flow rate equal to 680 veh/h. Lane width had a highly significant effect in both Model 1 and 365 Model 2. This variable was not evaluated in previous works because their PZs had the same lane 366 width.

Figure 4b shows the effect(s) of absolute vertical grade on P_{PZ} at traffic flow rate levels of 200, 400, 600, and 800 veh/h. Unlike previous studies (Mwesige, et al., 2016, Moreno, et al., 2013, Hegeman, 2008) where sites were located on flat terrain, in this study the absolute vertical grade had a wide range from 0.58 to 9.5%. The results show that absolute vertical grade had a positive impact on the passing rate, and the impact had a peak respect to the traffic flow rate. The vertical grade reduces the speed of heavy vehicles, which causes the platoons. To consider this impact, HCM used an adjusted factor to calculate the demand flow rate (TRB, 2010).

Figure 4c shows the rise in the passing rate in the subject direction (P_{PZ}) caused by an increase in the proportion of heavy vehicles in the subject direction (P_{HVS}) . Mwesige, et al. (2016) used the proportion of heavy vehicles in both directions (P_{HV}) ; however, this variable did not have a significant effect on the P_{PZ} . As the percentage of heavy vehicles in the subject direction increases, the number of platoons and their length grow, which increases the frustration and desire 379 of drivers behind to pass (Penmetsa, et al., 2015, Polus and Cohen, 2009). Hence, P_{HVS} better 380 explains the effect of the proportion of heavy vehicles on the passing rate in the subject direction 381 than the P_{HV} . By increasing the P_{HVS} , the platooning and PTSF will increase, which leads to an 382 increase in the demand for passing maneuvers. Drivers caught in a platoon behind slow vehicles 383 (heavy vehicles for example) will attempt to conduct more passing maneuvers. The ability to pass 384 is provided by the PZs if and when drivers find a gap in the opposite direction. Mwesige, et al. 385 (2016) showed that the passing rate increases with the P_{HV} to a peak at 35% and then decreases. In 386 this study, the highest levels of P_{HV} and P_{HVS} are 30.2 and 37.9%, respectively, and the results 387 show an increase in the passing rate with an increase in P_{HVS} consistently with Mwesige, et al. 388 (2016). The P_{HVS} did not have a significant effect on the number of passing maneuvers that start 389 and end in the NPZ (P_{PPZ}). It is for this reason that, as observed in the recorded videos, the platoon 390 discharges before the PZs, hence, most of the vehicles in the platoon start passing maneuvers in 391 the NPZ.

Figure 4d illustrates the effect of the proportion of motorcycles in the subject direction (P_{MCS}) on the P_{PZ} . The figure indicates that the P_{PZ} increased as the proportion of motorcycles in the subject direction increased. The reason for the increase in the passing rate is that the vehicles needed shorter gaps to pass. Furthermore, motorcycles had lower speed (51 km/h on average) and width and drove close to the road edge (for roads without shoulders) or to the shoulders, which made it easier for drivers to pass them.

It should be noted that high-speed motorcycles are forbidden in Iran. Hence, this variable might prove to be insignificant in other countries. In fact, in the study conducted by Hegeman (2008), motorcycles accounted for less than 2% of traffic flow, and they did not consider this variable in their model. Similarly, other previous studies did not evaluate the effect of the 402 proportion of motorcycles on the passing rate. Figure 5 illustrates the graphs of the passing rate in 403 the subject direction and the traffic flow rate in both directions. Figures 4 was produced for V_G = 404 3.5%, P_{HVS} = 30%, L_W = 3.5 m, L_{PZ} = 300 m and P_{MCS} = 0%. These values were selected so as to 405 be consistent with studies conducted by Mwesige, et al. (2016) and Moreno, et al. (2013) and make 406 it possible for comparing. The graphs of P_{PZ} were plotted at five directional split levels of 20, 50, 407 60, 70, and 80%.

408 The results show that the directional split has a significant effect on the passing rate as shown by 409 the fact that an increase in the directional split in the subject direction results in a similar increase 410 in the passing rate. The directional split also had a significant effect on the studies of Mwesige, 411 et al. (2016) and Moreno, et al. (2013). Hegeman (2008) employed both subject and opposite 412 traffic volumes in her model instead of using two-way traffic volume and directional splits. The 413 graphs of previous studies (Mwesige, et al., 2016, Moreno, et al., 2013, Hegeman, 2008) and 414 P_{PPZ} in Figure 5 were plotted for a directional split of 50%. The model of Mwesige, et al. (2016) 415 (Equation (5)) is very close to the model that predicts the P_{PZ} (Equation 14). Both models 416 predicted the passing rate of passing maneuvers that ended in the PZs and also employed most of 417 the explanatory variables for estimation compared to other previous studies, and it could be the 418 reason for the more accurate prediction of these two models. Another reason could be the 419 similarity in the traffic culture of the countries.

The models proposed by Hegeman (2008) and Moreno, et al. (2013) (Equation (3) and (4) underestimated the passing rate. Their studies were carried out in developed countries, which might explain the significant differences between passing rate values in their work and that of Mwesige, et al. (2016). Another reason for the lower prediction by Hegeman (2008) is due to the *Dutch Sustainable Safety Program*, which recommended that authorities design two-lane highways along which passing is prohibited, even where there is sufficient sight distance. In the
study of Moreno, et al. (2013), the two studied PZs were too close to each other (less than 350 m),
hence a reduction in the passing rate was possible.

Figure 5 shows that the passing rate of maneuvers ending in the PZ (P_{PZ}) is nearly two times that of the passing rate of maneuvers that started and ended in the PZ (P_{PPZ}). Furthermore, the PZ length had an insignificant effect on P_{PZ} . All of which implies that increasing the length of the short PZs does not lead to a significant increase in the passing rate and, by extension, more fluid traffic operations.

However, the high passing rates observed demonstrate that short PZs irrespective of their length have a significant effect on the operation of highways in Iran. Harwood, et al. (2008) observed that only 0.4% of all vehicles passed in the short PZs, while this value was 5% on average in Iran. Moreno, et al. (2016) studied the effects of PZ length (250-5000m) on the operational performance of two-lane highways. Their findings showed that average PZ length had an effect on traffic performance, while short PZs (250m) did not.

439 As shown in Figure 5, the passing rate for maneuvers ended in the PZ (P_{PZ}) peaked at the flow

440 rate of 680 veh/h, after this a reduction was observed. This maximum value represents the

441 passing capacity, which indicates the maximum effectiveness of PZs respect to flow rate in both442 directions.

Figure 5 shows the passing capacity for maneuvers that start and end in the PZ (P_{PPZ}) at the flow rate of 597 veh/h. Moreno, et al. (2013) found that the peak passing rate occurred at 600-700 veh/h, which is similar to the findings of this research. However, in the study conducted by Mwesige, et al. (2016), the passing rate reached its capacity at the flow rate in both directions at 400 veh/h. The maximum flow rate that they observed in their study was 426 veh/h. Hence, they could not estimate the effect of superior levels of flow rate on the passing rate. They stated
that it was not possible to observe the passing capacity at 600 veh/h because the average
headways were too short (Mwesige, et al., 2016). However, it should be noted that by increasing

the flow rate, the platoons will increase and, commonly, several drivers in the platoons use a

452 single gap in the opposite direction to pass the slow vehicle together.

453

454 **6.** Conclusion

This study evaluated the effects of geometric characteristics and traffic-related explanatory variables on the passing rate of short passing zones (PZs). The effectiveness of variables was obtained by generating Poisson regression models from aerial data collected from seven PZs.

458 The main conclusions of the study were as follows:

- the length of the short PZs did not have a significant effect on the passing rate of
 maneuvers ending in the PZ; hence, limited increases in length of the short PZs would not
 help to improve traffic operations since, in practice, a significant proportion of drivers start
 their passing maneuvers before reaching short PZs. Nevertheless, the results showed that
 short PZs had an important role in improving traffic operations in Iran irrespective of their
 length;
- the lane width had a highly significant effect on the passing rate;
- by increasing the proportion of motorcycles in the subject direction, the passing rate in the
 subject direction witnessed a significant increase;
- the proportion of heavy vehicles in the subject direction was a significant variable when
 estimating the passing rate in the subject direction, while the proportion of heavy vehicles
 in both directions proved insignificant;

472

• the passing capacity of the PZs occurred at the flow rate of 680 veh/h, and the maximum increase rate in the passing rate occurred at the flow rate of 400 veh/h.

473 The models presented support planners and designers in developing safety and operational 474 analyses. Traffic flow rate is an important variable in planning for a highway. Based on the 475 projected figures for passing capacity and predicted flow rate, planners could select between 476 two-lane highways and other types of infrastructure. Limited increments in length would not lead 477 to an improvement in traffic operations at short PZs. Other geometrics characteristics, e.g., the 478 lane width, could play a more significant role. However, the increment in length of short PZs could 479 significantly improve safety (Moreno, et al. (2015). Accordingly, future research should 480 investigate the effect of short PZs on safety. From a safety perspective, increasing the flow rate 481 beyond the passing capacity serves only to decrease the passing rate and increase driver frustration 482 levels, which in turn could result in dangerous passing maneuvers. Furthermore, since in Iran, 483 motorcycles travel at a slower speed than passenger cars, the estimated models suggest that an 484 increase in the proportion of motorcycles increases the passing rate. Hence, safety experts should 485 base their evaluation of safety performance on traffic volumes, passing capacity, and the 486 proportion of motorcycles.

Heavy vehicles travel at a slower speed than passenger cars, and their speed decreases further as the absolute vertical grade increases. The result of this work showed that an increment in the proportion of heavy vehicles and/or the absolute vertical grade increases the passing rate. However, the passing maneuver is a demanding and dangerous task, so planners and designers should consider the safety implications of the proportion of heavy vehicles and vertical grades when deciding between two-lane highways and other facility types. To achieve more homogenous speeds, lower vertical grades have to be assumed at the design stage if a high percentage of heavyvehicles is expected.

495 Some variables such as shoulder width, pavement surface quality, and weather conditions 496 need to be evaluated in future studies. Llorca, et al. (2013) concluded that the passing driver's 497 behavior was different at night. Hence, an investigation into the passing rate at nighttime is needed. 498 The posted speed limit could also affect the passing rate. In this study, there was not enough 499 variation in this variable for analysis. Moreover, the strategies to enforce the speed limit could be 500 important, and this aspect also requires field study in PZs. A simulation study conducted by Ghods 501 and Saccomanno (2016) showed that differential speed strategies had a significant effect on the 502 passing rate of passenger cars when overtaking heavy vehicles.

503 This study focused on short PZs, with length values between 164 and 345 m (Table 1). A 504 wider range of PZ lengths could help to find cases where the length of PZ has a significant effect 505 on the passing rate. To generalize the proposed models across different countries, future studies 506 need to be carried out.

507

508 Data Availability Statement

509 Some or all data, models, or code that support the findings of this study are available from the 510 corresponding author upon reasonable request.

511

512 Acknowledgment

513 This research was funded by Tarbiat Modares University.

514 Notation list

515 D_S = directional split of traffic volume

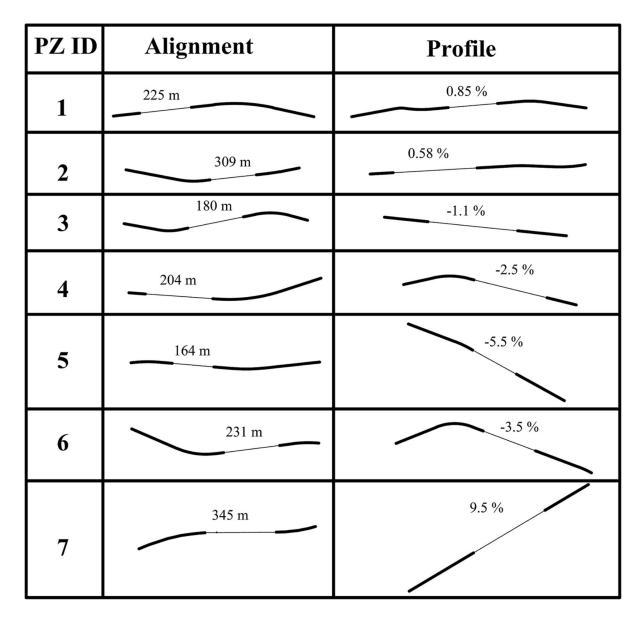
- L_{PZ} = passing zone length
- $L_W = \text{lane width}$
- N_{PPZ} = number of passing maneuvers per 15 minutes starting and ending in PZ in subject
- 519 direction
- 520 NPZ = no-passing zone
- N_{PZ} = number of passing maneuvers per 15 minutes ending in the passing zone (PZ) in the
- 522 subject direction
- P_{HV} = proportion of heavy vehicles in both directions
- P_{HVS} = proportion of heavy vehicles in the subject direction
- P_{MCS} = proportion of motorcycles in the subject direction
- P_{PPZ} = number of passing maneuvers per hour starting and ending in PZ in subject direction
- P_{PZ} = number of passing maneuvers per hour ending in the passing zone (PZ) in the subject
- 528 direction
- 529 PTSF = percent time spent following
- PZ = passing zone
- V = flow rate in both directions
- V_G = absolute vertical grade
- $V_O =$ flow rate in the opposite direction
- V_S = flow rate in the subject direction

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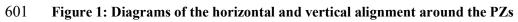
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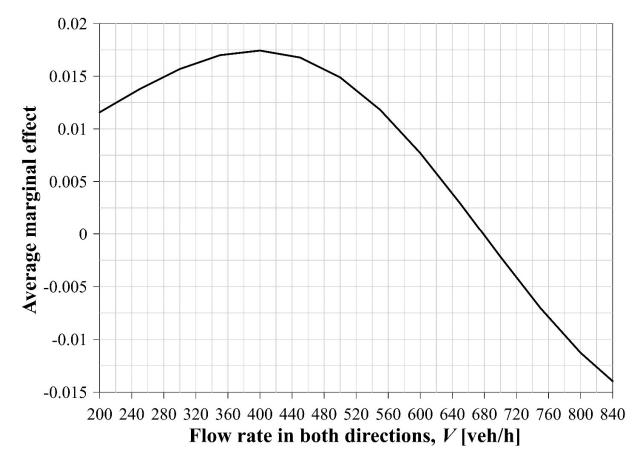




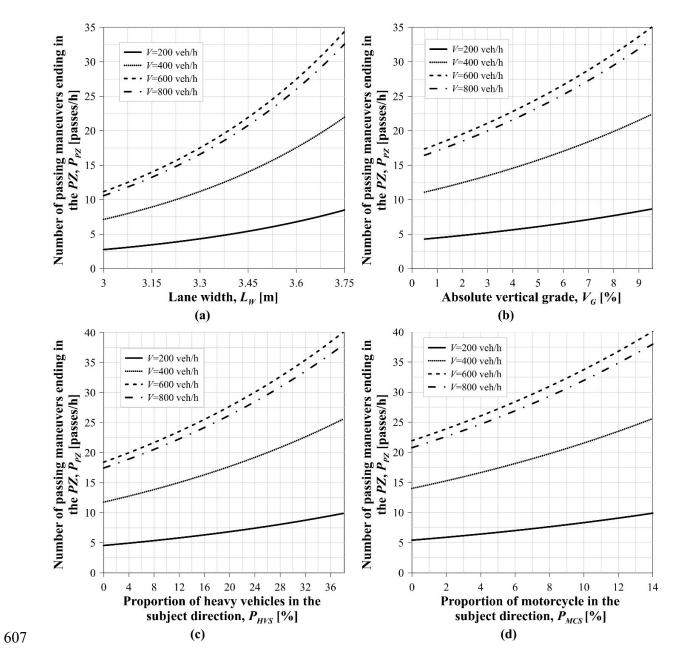




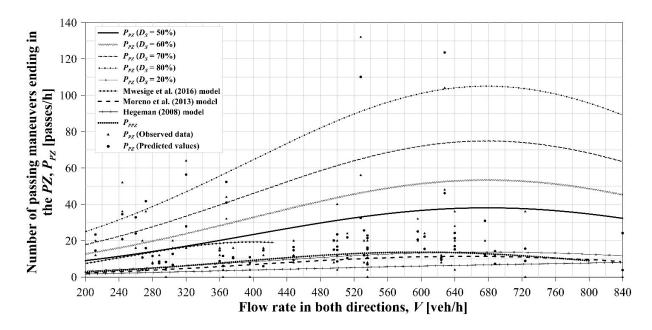
603 Figure 2: Phantom 4 Pro drone



606 Figure 3: The average marginal effects of the flow rate in both directions on NPZ



608 Figure 4: Effects of explanatory variables of Model 1 on the *P_{PZ}*



609

610 Figure 5: Effect of the flow rate in both directions on the passing rate and comparison of this

611 study's models and previous works

613 Table 1: List of existing studies and their corresponding contributions (e.g., explanatory variables)

614 and shortcomings

						Ε	xplanatory	variab	les					
Studies	V_S	V_O	V	$\overline{\mu}_{s}$	σ_{s}	P(g>30)	P(road)	L_{PZ}	D_S	V_G	S ₈₅	P_{HV}	k	Shortcomings
Wardrop (1952)	*			*	*									A theoretical mode
Daganzo (1975)			*											A theoretical mode
McLean (1989)						*	*							A theoretical mode
Dommerholt and Botma (1988)	*	*	*		*								*	A theoretical mode
Tuessiner and Enhance (2006)	*													Using a linear
Tuovinen and Enberg (2006)														regression mode
Hageman (2008)	*	*												Using a linear
Hegeman (2008)														regression mode
Moreno et al. (2013)			*					*	*					
Mwesige et al. (2016)			*					*	*	*	*	*		
V_S = traffic flow in the subject dir	rection	ı [veh/l	1]					P(roa	<i>d)</i> = pi	oporti	on of 1	oad ler	ngth s	uitable for passing
$V_o =$ traffic flow in the opposite c	lirectio	on [veł	ı/h]					$L_{PZ} =$	length	of PZ	[<i>m</i>]			
V = traffic flow in both direction	[veh/h	ı]						$D_S = 0$	directio	onal sp	lit in t	he subj	ect di	irection [%]
$\overline{\mu}_s$ = average of space mean spee	ds [kn	n/h]						$V_G = i$	absolut	e verti	ical gra	ade [%]		
standard deviation of succes		لممعه	[]/le	1				$S_{85} = 8$	85 th per	rcentil	e speed	d [km/h]	
σ_s = standard deviation of space	mean	speed	[KIII/N]				$P_{HV} =$	propo	rtion o	f heav	y vehic	les [9	%]
P(g>30) = proportion of time that	t there	e is a ga	ip larg	er thar	the cr	itical gap (30) s)	k = de	ensity o	of vehi	cles [p	c/km]		

PZ ID	Highway (*)	Length	Width	Shoulder width	Absolute vertical grade	Posted speed limit	Passing maneuvers recorded	15-min period observations
		[m]	[m]	[m]	[%]	[Km/h]	[#]	[#]
1	JF	225	3.45	-	0.85	85	73	16
2	JF	309	3.45	-	0.58	85	58	14
3	JF	180	3.25	-	1.1	85	24	10
4	JF	204	3.25	-	2.5	85	17	8
5	JB	164	3.00	-	5.5	95	15	4
6	JB	231	3.00	-	3.5	95	18	8
7	JK	345	3.75	1.2	9.5	85	174	16

Table 2: Geometric characteristics of the PZs and total number of observed passing maneuvers at

617 each site for 15-min periods in each direction

618 (*) JF = Jiroft-Faryab, JB = Jiroft-Baft, and JK = Jiroft-Kerman

619

	М	NC.	M	Confidence
Variable (Unit)	Mean	Min.	Max.	interval (95%)
Number of passing maneuvers ending in PZ in the subject	5.0	0	22	
direction, N _{PZ} (Passes per 15 min)	5.0	0	33	(3.8, 6.2)
Number of passing maneuvers starting and ending in PZ in	2.3	0	21	(1521)
subject direction, N _{PPZ} (Passes per 15 min)	2.3	0	21	(1.5, 3.1)
PZ length, L_{PZ} (m)	254.60	164	345	(240.2, 269.0)
Upstream NPZ length, L_{UNPZ} (m)	1265.5	130	5010	(946.4, 1584.6)
Lane width, $L_W(\mathbf{m})$	3.39	3.00	3.75	(3.34, 3.45)
Absolute vertical grade, V_G (%)	3.4	0.6	9.5	(2.6, 4.1)
Flow rate in both directions, V (veh/h)	470.4	212	840	(434.6, 506.3)
Flow rate in the subject direction, V _S (veh/h)	235.2	80	648	(212.7, 257.8)
Flow rate in the opposite direction, V ₀ (veh/h)	235.2	80	648	(212.7, 257.8)
Directional split of traffic volume, D_S (%)	50	22.9	77.1	(47.6, 52.4)
Proportion of heavy vehicles in both directions, P_{HV} (%)	8.2	0	30.2	(6.4, 10.1)
Proportion of heavy vehicles in the subject direction, P_{HVS} (%)	8.6	0	37.9	(6.4, 10.8)
Proportion of motorcycle in the subject direction, P_{MCS} (%)	3.0	0	13.9	(2.2, 3.7)
Mean free-flow speed, M_{FFS} (km/h)	85.3	63.7	102.0	(83.7, 87.0)
Standard deviation of free-flow speed, σ_s (km/h)	16.0	9.5	27.6	(15.1, 16.9)
85 th Percentile free-flow speed, S_{85} (km/h)	101.3	71.0	129.4	(99.1, 103.4)

620 Table 3: Descriptive statistics of explanatory variables

variables	Mode	11	Moo	del 2
variables	β-Estimate	Z-value	β -Estimate	Z-value
L _{PZ}	-	-	0.00786	2.69 ^b
L_W	1.4989	3.82ª	1.71082	2.08°
V_G	0.0779	2.99 ^b	0.06994	2.04 ^c
V	0.00852	3.39 ^b	0.01218	3.00 ^b
V^2	-6.28×10 ⁻⁶	-2.41°	-10.2×10 ⁻⁶	-2.31°
D_S	0.03381	6.19ª	0.02636	4.55 ^a
P _{HVS}	0.02044	2.00 ^c	-	-
P_{MCS}	0.04307	1.78 ^d	-	-
constant	-8.4586	-7.44ª	-12.31318	-5.20ª
Test	χ^2	<i>p</i> -value	χ^2	<i>p</i> -value
Overall model evaluation				
Likelihood ratio test	226.73	0.0000	208.69	0.0000
Goodness-of-fit				
Deviance test	86.925	0.061	90.559	0.042
Pearson test	74.03	0.288	84.488	0.099
	$\overline{\chi}^2$	<i>p</i> -value	$\overline{\chi}^2$	<i>p</i> -value
Overdispersion				
LR test of α	0.00000	0.500	0.09	0.379
Random-effects				
LR test of ϕ	0.00	1.00	0.00	1.00
Cragg-Uhler R ²	0.95	0	0.93	9
McFadden's R ²	0.41	8	0.47	9
Sample size	76		76	

622 Table 4: Estimated Poisson model parameters (Model 1 predicts N_{PZ}, Model 2 predicts N_{PPZ})

623 (a) significance level at 0.001. (b) significance level at 0.01. (c) significance level at 0.05. (d) significance level at 0.1.

Variables	Average marginal effects	Confidence interval (95%)
L_W	7.47	(3.57, 11.38)
V_G	0.39	(0.13, 0.65)
D_S	0.17	(0.11, 0.22)
P_{HVS}	0.10	(0.001, 0.20)
P_{MCS}	0.21	(-0.022, 0.45)

625 Table 5: Average marginal effects of explanatory variables on the number of passing maneuvers

ending in PZ per 15 minutes [passes/15-min/one-unit change in variable]

627

628 Table 6: Average marginal effects of explanatory variables on the number of passing maneuvers

Variables	Average marginal effects	Confidence interval (95%)		
L_{PZ}	0.018	(0.004, 0.031)		
L_W	3.92	(0.17, 7.66)		
V_G	0.16	(0.004, 0.316)		
D_S	0.06	(0.033, 0.088)		

629 that start and end in PZ per 15 minutes [passes/15-min/one-unit change in variable]