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A Perturbation Scheme for Passivity Verification and Enforcement of Parameterized Macromodels

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Abstract—This paper presents an algorithm for checking and enforcing passivity of behavioral reduced-order macromodels of LTI systems, whose frequency-domain (scattering) responses depend on external parameters. Such models, which are typically extracted from sampled input-output responses obtained from numerical solution of first-principle physical models, usually expressed as Partial Differential Equations, prove extremely useful in design flows, since they allow optimization, what-if or sensitivity analyses, and design centering.

Starting from an implicit parameterization of both poles and residues of the model, as resulting from well-known model identification schemes based on the Generalized Sanathanan-Koerner iteration, we construct a parameter-dependent Skew-Hamiltonian/Hamiltonian matrix pencil. The iterative extraction of purely imaginary eigenvalues of the pencil, combined with an adaptive sampling scheme in the parameter space, is able to identify all regions in the frequency-parameter plane where local passivity violations occur. Then, a singular value perturbation scheme is setup to iteratively correct the model coefficients, until all local passivity violations are eliminated. The final result is a corrected model, which is uniformly passive throughout the parameter range. Several numerical examples demonstrate the effectiveness of the proposed approach.

Index Terms—Macromodeling, parameterized modeling, transmission lines, rational approximation, reduced order modeling, Hamiltonian matrices, singular values, perturbation.

I. INTRODUCTION AND MOTIVATION

This paper addresses the general problems of passivity verification and passivity enforcement for behavioral models of LTI systems, whose transfer function depends on external parameters. This problem arises in all those application areas that require reduced-order, compact macromodels that can be simulated efficiently either in frequency or time-domain using standard Ordinary Differential Equation (ODE) or legacy circuit solvers of the SPICE class, in order to assess the performance of a given underlying complex circuit or system [1], [2], [5], [6]. Typical application scenarios include Signal and Power Integrity verification of digital, analog, mixed-signal and RF systems including all parasitic effects of electrical interconnects at chip, package and board level, reduced-order modeling of circuit blocks for speeding up simulation tasks, or Computer Aided Design via virtual prototyping, including sensitivity, what-if and design centering processes.

Behavioral reduced-order macromodels are typically derived either through model order reduction of large circuit equations cast in state-space or descriptor form [3], [4], [6], whenever such characterizations are available (e.g., through a partial element equivalent circuit extraction, or direct spatial discretization of Maxwell’s equations), or through a model identification from measured or simulated port responses [1], [7]–[11]. Only the latter case is considered in this work, due to its widespread adoption by industry, academia, and simulation tool vendors. Therefore, we target here the construction of a passive model in state-space or descriptor form of a generic LTI multiport system starting from sampled frequency responses at its terminals.

Several approaches are available for macromodel generation in the univariate case, where the system transfer function or matrix depends only on frequency (see [1] for a comprehensive review). Since model extraction is based on fitting a prescribed model structure to available data through some numerical optimization, the macromodel responses are inevitably affected by some approximation error. This error cannot be avoided, since most often the underlying system is characterized by non-rational transfer functions (e.g., in the common case of distributed systems with lossy and dispersive materials). Fitting a (finite-order) rational model to represent such systems cannot lead to exact identification. Such approximation error can be sufficient to make a model non-passive, even when it is supposed to represent a passive physical system. Lack of passivity is very critical and may lead to instabilities when using the models during system-level simulations [14], [15]. Therefore, any passivity violation must be carefully identified and removed.

Passivity verification and enforcement of univariate models is well-established. The common tool that provides an algebraic test for passivity is the so-called Hamiltonian matrix associated to a state-space realization of the model [16], [23]. The presence of purely imaginary eigenvalues of such matrix pinpoints passivity violations (e.g., frequency bands where the model has active behavior). In case of such violations, several model perturbation methods can be used to correct the model coefficients and recover model passivity [1], [26]. Most prominent methods are based on Hamiltonian eigenvalue perturbation [22], [23], singular value/eigenvalue perturbation [24], [25], Positive or Bounded Real Lemma [17] enforcement based on semidefinite programming [27], \( H_\infty \) norm optimization via non-smooth (convex) optimization [28] or localization methods [29].

The main objective of this work is the extension of the above passivity verification and enforcement methods to the multivariate case, where the transfer function of the model depends not only on frequency, but also on additional external parameters. Examples can be geometrical or material properties of the underlying physical circuit or system [30], [31], [33], [34], [40], temperature or biasing conditions [35], etc.
Few approaches are able to provide parameterized models that are uniformly passive throughout the parameter domain. These approaches impose some constraints on the model structure and/or the model identification procedure. For instance, we can cite the interpolatory approaches that first construct univariate models for fixed parameter values, and then recover a multivariate model by interpolating such “grid” or “root” macromodels. If the root macromodels are passive and if a passivity-preserving interpolation scheme is used, also the multivariate model will be uniformly passive [37]–[40]. This approach has two main problems. First, passivity-preserving interpolation schemes are overconstrained and may lead to inaccurate behavior of the interpolant. This problem is usually addressed by increasing the number of root macromodels, thus degrading the efficiency of the identification process. Second, the multivariate model is cast as a linear combination of root macromodels. Therefore, if each root macromodel is characterized by a given order \( n \), the multivariate model will have a larger order \( L \), where \( L \) is the number of root macromodels contributing to the interpolation. On one hand, this model structure may not be related to the physics of the underlying system (if \( n \) natural frequencies are required, this should be true for all parameter values). On the other hand, model complexity is larger than strictly necessary.

The approach that we follow in this work is different. We embed the parameter dependence through a dedicated model structure, based on a suitable expansion into frequency-dependent and parameter-dependent basis functions [32]–[35]. Thus, model structure is general and can be tuned by selecting the appropriate basis functions for the specific application at hand. Model coefficients are then identified through a Generalized Sanathanan-Koerner iteration [41]. This approach is standard [30], [32].

The novel contribution of this work is twofold. First, we devise an adaptive sampling scheme in the parameter space based on some special features of the Hamiltonian eigenspectrum. This scheme is able to pinpoint all passivity violations throughout the frequency and parameter plane. Second, we extend a known passivity enforcement scheme for univariate models based on singular value perturbation to the multivariate case. The result is an iterative passivity enforcement scheme for multivariate macromodels that is able to guarantee uniform passivity as well as model accuracy with constant model order throughout the parameter space. To the best of Authors’ knowledge, this is the first passivity verification and enforcement scheme that is applicable to multivariate models. We focus this work on the case of a single external parameter, or equivalently on bivariate macromodels. The extension to the multidimensional case will be documented in a future report.

II. PRELIMINARIES AND NOTATION

Let us consider a generic LTI multiport structure with transfer matrix \( \tilde{H}(s; \theta) \), where \( s \) is the Laplace variable, and \( \theta \in \Theta = [\theta_{\min}, \theta_{\max}] \) is some external parameter, e.g., some material or geometrical characteristic of the underlying physical system. Throughout this work, it is assumed that \( \tilde{H}(s; \theta) \in \mathbb{C}^{P \times P} \) is the scattering matrix of the system at well-defined ports. The accent \( \tilde{\cdot} \) will be used to label the original or “true” system responses, that are generally unknown.

The “true” response is assumed to be available, either via direct measurements or via numerical calculation based on a first-principle model (e.g., the numerical solution of Maxwell’s equations for an electromagnetic system) at a set of fixed frequencies \( f_k, k = 1, \ldots, \tilde{k} \) over a given frequency band \( [f_{\min}, f_{\max}] \), with \( f_1 = f_{\min} \) and \( f_{\tilde{k}} = f_{\max} \), and for a set of fixed parameter values \( \theta_m, m = 1, \ldots, \tilde{m} \) spanning the range \( [\theta_{\min}, \theta_{\max}] \). We will denote this characterization as

\[
\tilde{H}_{k,m} = \tilde{H}(j2\pi f_k; \theta_m), \quad k = 1, \ldots, \tilde{k}, \quad m = 1, \ldots, \tilde{m}.
\]

Since a closed-form dependence of the true system response on \( s \) and \( \theta \) is not available, there comes the need of constructing a parameterized model that fits or interpolates the above data points, and that can be solved efficiently in frequency and time domain. We assume here the following model structure

\[
H(s; \theta) = \frac{\mathbf{N}(s, \theta)}{\mathbf{D}(s, \theta)} = \sum_{n=0}^{\tilde{n}} \mathbf{R}_n(\theta) \varphi_n(s),
\]

where the frequency-dependent basis functions \( \varphi_n(s) \) are partial fractions based on a prescribed set of distinct \( \bar{n}_r \) real poles \( q_n \in \mathbb{R}^- \) and \( \bar{n}_c \) complex pole pairs \( p_n = p_n^r \pm j p_n^\ell \in \mathbb{C}^- \), where \( \varphi_0(s) = 1 \),

\[
\varphi_n(s) = (s - q_n)^{-1}, \quad \text{for} \quad n = 1, \ldots, \bar{n}_r,
\]

and

\[
\varphi_{n, \nu+2\nu}(s) = (s - p_{\nu}^{-1})^{-1} + (s - p_{\nu}^{\ell})^{-1}
\]

\[
\varphi_{n, \nu+2\nu}(s) = j(s - p_{\nu}^{-1})^{-1} - j(s - p_{\nu}^{\ell})^{-1}
\]

for \( \nu = 1, \ldots, \bar{n}_c \), where \( ^* \) denotes the complex conjugate. The (real-valued) numerator and denominator coefficients are further expressed as

\[
\mathbf{R}_n(\theta) = \sum_{\ell=1}^{\tilde{\ell}} \mathbf{R}_{n,\ell} \xi_\ell(\theta), \quad r_n(\theta) = \sum_{\ell=1}^{\tilde{\ell}} r_{n,\ell} \xi_\ell(\theta)
\]

where the basis functions \( \xi_\ell(\theta) \) are responsible for reproducing the variations induced by the external parameter \( \theta \). We remark that the adopted model structure is the same as discussed in [1] and originally introduced in [32], [36], [37], [39], [40], where polynomials, piecewise polynomials, or trigonometric polynomials were used as basis functions \( \xi_\ell \). This work makes no a-priori assumption on the specific choice of basis functions, which should be defined considering the particular application at hand.

Model structure (2) guarantees that \( H(s; \theta) \) is a rational function of \( s \), with implicitly parameterized poles and residues (explicit parameterization of model poles is to be avoided, due to the possibly non-smooth behavior in case of bifurcations [36]). Thanks to the assumed linear dependence of both model numerator \( \mathbf{N}(s, \theta) \) and denominator \( \mathbf{D}(s, \theta) \) on the respective coefficients \( \mathbf{R}_{n,\ell} \) and \( r_{n,\ell} \), the latter can be easily evaluated during a model identification step by enforcing the fitting condition

\[
H(j2\pi f_k; \theta_m) \approx \tilde{H}_{k,m}, \quad k = 1, \ldots, \tilde{k}, \quad m = 1, \ldots, \tilde{m}.
\]

In least squares sense. This is achieved here through a linear relaxation of (6) known as (Generalized) Sanathanan-Koerner
(GSK) iteration [1], [32], [41]. This approach transforms the nonlinear least-squares problem arising from (6) into an iterative sequence of weighted linear least squares problem, whose solution is straightforward. This procedure is standard in parameterized model identification and is not further discussed here. The Reader is referred to [1] for more details on the GSK iteration, and to [44], [45] for a discussion on its convergence properties. We finally remark that, in order to guarantee uniqueness in the model representation, we normalize the model coefficients by setting $r_{0,1} = 1$.

III. PROBLEM STATEMENT

The above-mentioned GSK iteration is not able to enforce model and passivity by construction, since no explicit passivity constraints are enforced during model identification. We recall that the model is passive for a given parameter value $\vartheta$ if and only if the corresponding scattering matrix $H(s; \vartheta)$ is Bounded Real [12], [13]:

1. $H(s; \vartheta)$ is regular for $\Re\{s\} > 0$,
2. $H^*(s; \vartheta) = H(s^*; \vartheta)$,
3. $I_P - H^H(s; \vartheta)H(s; \vartheta) \geq 0$ for $\Re\{s\} > 0$,

where $H$ is the Hermitian transpose, and $I_P$ is the identity matrix of size $P$. Condition 1) is related to (asymptotic) stability, which is here assumed a priori (as can be readily verified by a suitable parameter sweep of the parameter-dependent model poles), whereas condition 2) ensuring a real impulse response is automatically satisfied thanks to the assumed model structure (2), Condition 3), implying no energy gain from the model throughout the open right complex plane, is instead more difficult to check and to enforce. Note that it is sufficient to check condition 3) only on the imaginary axis $s = j\omega$,

$$I_P - H^H(j\omega; \vartheta)H(j\omega; \vartheta) \geq 0 \quad \forall \omega \in \mathbb{R},$$

which in turn is equivalent to

$$\sigma_{\max}\{H(j\omega; \vartheta)\} \leq 1, \quad \forall \omega \in \mathbb{R},$$

where $\sigma_{\max}\{\cdot\}$ extracts the largest singular value of its matrix argument.

Several solutions are in fact available for checking and enforcing (8) on univariate models that depend only on frequency, which in our setting would correspond to the model instantiated for a fixed value of the external parameter $\vartheta$. For a comprehensive review of such methods, the Reader is referred to [1] and to [26]. What we are interested in this work is the uniform passivity of the parameterized model $H(s; \vartheta)$ throughout the parameter range, so that the above Bounded Realness conditions, in particular (8), will hold $\forall \vartheta \in [\vartheta_{\min}, \vartheta_{\max}]$.

The approach that is pursued here can be regarded as an extension to the multivariate case, of existing standard approaches valid for non-parameterized systems. We start from some initial non-passive model $H(s; \vartheta)$, and we perturb its numerator coefficients as

$$\hat{R}_{n,\ell} = R_{n,\ell} + \Delta R_{n,\ell} \quad \forall n, \ell$$

by determining the correction $\Delta R_{n,\ell}$ that is required to ensure that the perturbed scattering response of the parameterized model $H(s; \vartheta)$ is uniformly Bounded Real in the prescribed parameter range. Note that we do not perturb the denominator coefficients, since we would like to preserve the model poles, i.e., the zeros of the denominator $D(s, \vartheta)$. This choice is standard in practically all passivity enforcement methods applicable to univariate models.

Setting up the model perturbation requires a precise localization of the regions in the frequency-parameter plane where the Bounded Realness condition (8) is violated. This problem is addressed in Section V, which first casts some known results for univariate models in the parameterized framework (Section V-A) and then extends the check to the multivariate case (Section V-B). Section VI presents the main passivity enforcement scheme. In particular, Section VI-A constructs a cost function whose minimization will ensure preservation of model accuracy during perturbation; and Section VI-C presents the main iterative passivity enforcement scheme. All these developments require the construction of a parameter-dependent descriptor realization of the model, which is discussed next.

IV. DESCRIPTOR REALIZATIONS

Starting from model (2), we define the following descriptor realization

$$\begin{align*}
E \dot{x} &= A(\vartheta)x + Bu \\
\dot{y} &= C(\vartheta)x
\end{align*}$$

where $x \in \mathbb{R}^{N+P}$ denotes internal generalized states with $N = \bar{n}P$, and $u, y \in \mathbb{R}^p$ are incident and reflected scattering waves at the model ports. The descriptor matrices are constructed as

$$E = \begin{bmatrix} I_N & 0_{N,P} \\ 0_{P,N} & 0_{P,P} \end{bmatrix}, \quad A(\vartheta) = \begin{bmatrix} A_0 & B_0 \\ C_2(\vartheta) & D_2(\vartheta) \end{bmatrix},$$

$$C(\vartheta) = \begin{bmatrix} C_1(\vartheta) \\ D_1(\vartheta) \end{bmatrix}, \quad B(\vartheta) = \begin{bmatrix} 0_{N,P} \\ -I_P(\vartheta) \end{bmatrix},$$

where $0_{L,K}$ denotes the all-zero matrix block of size $J \times K$. In (11), the matrices $\{A_0, B_0, C_2(\vartheta), D_2(\vartheta)\}$ and $\{A_0, B_0, C_2(\vartheta), D_2(\vartheta)\}$ provide regular state-space realizations of model numerators $N(s, \vartheta)$ and “extended” denominator $D(s, \vartheta)I_P$, respectively. Individual matrices in these realizations are here defined as

$$A_0 = \text{blkdiag}(A_{0r}, A_{0c})$$

$$B_{0r} = \begin{bmatrix} B_{0r}^T \end{bmatrix}$$

where $T$ is the matrix transpose, with

$$A_{0r} = \text{blkdiag}(q_{0r}I_P)_{n=1}^{\bar{n}_r}$$

$$A_{0c} = \text{blkdiag}\left\{p_{n}^{\prime}I_P \begin{bmatrix} p_{n}^{\prime}I_P \\ \hat{p}_{n}^{\prime}I_P \end{bmatrix} \right\}_{n=1}^{\bar{n}_c}$$

$$B_{0r} = [1, \ldots, 1]^T \otimes I_P$$

$$B_{0c} = [2, 0, \ldots, 2, 0]^T \otimes I_P$$

where $\otimes$ represents the Kronecker product.
where ⊗ is the Kronecker product, and
\[
\begin{align*}
C_1(\vartheta) &= [R_1(\vartheta) \cdots R_\alpha(\vartheta)] \\
C_2(\vartheta) &= [I_P r_1(\vartheta) \cdots I_P r_\alpha(\vartheta)] \\
D_1(\vartheta) &= R_0(\vartheta) \\
D_2(\vartheta) &= I_P r_0(\vartheta)
\end{align*}
\]

It is straightforward to show that, with the above definitions, the transfer matrix associated to the descriptor form (10)
\[
H(s; \vartheta) = C(\vartheta)(sE - A(\vartheta))^{-1}B
\]
matches (2). We leave this tedious verification to the Reader, see also [32].

V. CHECKING PASSIVITY OF PARAMETERIZED MODELS

Based on the descriptor form (10) of the model transfer function, we can formulate an algorithm for checking its uniform passivity, with the objective of a precise localization of eventual passivity violations.

A. Passivity Check of Univariate Models

Let us consider the univariate model obtained from (15) by “freezing” the parameter \( \vartheta \). Checking the passivity of such a model is a standard problem [1]. The most effective tool to perform this check is the so-called Hamiltonian matrix associated to the model [16], [17], [23], which in the case of our descriptor form becomes a skew-Hamiltonian/Hamiltonian (SHH) matrix pencil, also denoted Generalized Hamiltonian Pencil [18]–[21]. We define the two block-matrices
\[
M(\vartheta) = \begin{bmatrix} A(\vartheta) & BB^T \\ -C^T(\vartheta)C(\vartheta) & -A^T(\vartheta) \end{bmatrix}, \quad K = \begin{bmatrix} E & 0 \\ 0 & E^T \end{bmatrix}
\]

where \( M(\vartheta) \) has Hamiltonian structure and \( K \) is skew-Hamiltonian, forming the SHH matrix pencil \((M(\vartheta), K)\). We have the following result:

**Theorem 1.** Assume that the pencil \((A(\vartheta), E)\) has no purely imaginary eigenvalues. Then, the following conditions are equivalent:

- \( \sigma = 1 \) is a singular value of \( H(j\omega_0; \vartheta) \) with \( |\omega_0| < +\infty \);
- \( j\omega_0 \) is an eigenvalue of the pencil \((M(\vartheta), K)\), i.e., there exist some vector \( v_0 \neq 0 \) such that
\[
M(\vartheta)v_0 = j\omega_0Kv_0
\]

The proof is straightforward and follows the same flow as the proof of Theorem 1 in [16]. See also [1], [19], [21].

Matrix \( K \) is singular, since \( E \) is singular, with \( \dim \ker(E) = P \) and \( \dim \ker(K) = 2P \). Therefore, we expect that the eigenspectrum of the pencil \((M(\vartheta), K)\) includes at least \( 2P \) infinite eigenvalues. We recall that infinite eigenvalues are characterized as pairs \((\alpha, \beta)\) such that
\[
\alpha M(\vartheta)v = \beta Kw
\]
for some vector \( v \), with \( \alpha = 0 \). These eigenvalues will be disregarded in the following, and we will consider only the finite eigenvalues of the pencil as
\[
\lambda = \beta/\alpha, \quad \alpha \neq 0
\]

that satisfy (17) for some \( v \neq 0 \). Of course, all these eigenvalues depend on the considered parameter value \( \vartheta \). Therefore, the (finite part of the) SHH eigenspectrum is parameter-dependent
\[
\Lambda(\vartheta) = \{ \lambda(\vartheta) = \beta(\vartheta)/\alpha(\vartheta) : \alpha(\vartheta) \neq 0 \}
\]

We remark that, due to the SHH structure of the pencil, the set \( \Lambda(\vartheta) \) for each \( \vartheta \) is symmetric with respect to both real and imaginary axis, so that eigenvalues are either occurring in pairs \( \{\lambda, -\lambda\} \) if they are purely real or purely imaginary, or quadruples \( \{\lambda, -\lambda, \lambda^*, -\lambda^*\} \) if they are complex-valued with a nonvanishing real and imaginary part [46]–[48].

Based on Theorem 1, if \( \Lambda(\vartheta) \) includes a purely imaginary eigenvalue \( j\omega_0 \), then the transfer matrix of the model has a singular value \( \sigma(j\omega_0, \vartheta) = 1 \). We conclude that the singular value trajectory for fixed \( \vartheta \) in a neighborhood of \( j\omega_0 \) crosses the unit threshold, which is the critical condition leading to a passivity violation, see (8). This violation will be for \( \sigma > \omega_0 \) if \( \sigma \) increases with frequency, and vice versa. Figure 1 illustrates the relationship between imaginary SHH eigenvalues and localized passivity violations based on singular value trajectories.

As proposed in [23], we define the two sets
\[
\chi(\vartheta) = \{ \omega_1(\vartheta) \}_{i=1}^{\nu(\vartheta)}, \quad \bar{\chi}(\vartheta) = \{ 0 \} \cup \chi(\vartheta) \cup \{ +\infty \}
\]

collecting the frequencies \( \omega_i(\vartheta) > 0 \) corresponding to the the \( \nu(\vartheta) \) purely imaginary SHH eigenvalues \( \lambda_i(\vartheta) = j\omega_i(\vartheta) \) with positive (finite) imaginary part, sorted in ascending order. The augmented set \( \bar{\chi}(\vartheta) \) includes also the DC point \( \omega_0 = 0 \) and the infinite frequency \( \omega = +\infty \). The elements of \( \bar{\chi}(\vartheta) \) thus induce a subdivision of the positive frequency axis \([0, +\infty]\) into \( \nu(\vartheta) + 1 \) disjoint subbands
\[
\Omega_i(\vartheta) = (\omega_i(\vartheta), \omega_{i+1}(\vartheta)), \quad i = 0, \ldots, \nu(\vartheta)
\]

where we defined \( \omega_0(\vartheta) = 0 \) and \( \omega_{\nu(\vartheta)+1}(\vartheta) = +\infty \). Since all intersections of some singular value trajectory with the threshold \( \sigma = 1 \) is captured in the set \( \chi(\vartheta) \), each subband \( \Omega_i(\vartheta) \) can be flagged either as locally passive (if all singular values are less than 1 in this band) or locally not passive (otherwise), by splitting the corresponding index sets into
\[
i \in \mathcal{I}_p(\vartheta) \text{ and } i \in \mathcal{I}_{np}(\vartheta), \text{ respectively.}
\]

**Fig. 1.** Graphical illustration of Theorem 1. Left panel: finite SHH eigenvalues (empty dots highlight purely imaginary eigs); right panel: singular value trajectories, which cross the unit threshold at frequencies \( \omega_i \), corresponding to purely imaginary SHH eigenvalues. Local singular value maxima \( \bar{\sigma}_i \) occur at frequencies \( \bar{\omega}_i \). For this example, \( P = 2, \nu = 2, \mathcal{I}_p = \{ 0, 2 \} \) and \( \mathcal{I}_{np} = \{ 1 \} \).
passivity violation for each non-passive band is defined as the maximum singular value $\tilde{\sigma}_i(\vartheta)$, occurring at some frequency $\tilde{\omega}_i(\vartheta) \in \Omega_i(\vartheta)$ with $i \in \mathcal{I}_{np}(\vartheta)$. Numerical estimates of such local maxima $(\tilde{\omega}_i(\vartheta), \tilde{\sigma}_i(\vartheta))$ are easily determined through local sampling. See Figure 1 for a graphical illustration.

B. Uniform Passivity Check of Parameterized Models

The procedure discussed in Section V-A allows a passivity check of univariate models throughout the frequency axis, without any need of sampling the singular values of the transfer function, but requiring only an algebraic determination of some SHH eigenvalues. In this Section, we enrich this test by extending its scope to a uniform passivity check throughout the parameter range $\vartheta \in \Theta$. The main tool that we consider is the auxiliary function $\psi(\vartheta)$, defined as

$$\psi(\vartheta) = \min_{\lambda(\vartheta) \in \Lambda(\vartheta)} \left| \frac{\text{Re} \{\lambda(\vartheta)\}}{\rho(\vartheta)} \right|$$

(22)

where $\rho(\vartheta) = \max_{\lambda(\vartheta) \in \Lambda(\vartheta)} |\lambda(\vartheta)|$ is the spectral radius of the SHH pencil (16), computed considering only the finite eigenvalues. This function vanishes when the set $\chi(\vartheta)$ is non-empty, denoting the presence of purely imaginary eigenvalues. Conversely, if $\chi(\vartheta)$ is empty and there are no imaginary eigenvalues, so that the univariate model is passive for the considered value of $\vartheta$, the function $\psi(\vartheta)$ measures the distance of the SHH eigenspectrum from the imaginary axis, normalized to the largest eigenvalue magnitude. This normalization makes the test less sensitive to the finite numerical precision in the eigenvalue computation, as well as independent on the units and/or normalizations used for the frequency variable.

The proposed check performs an adaptive sampling of $\psi(\vartheta)$ within the range $\vartheta \in \Theta = [\vartheta_{\text{min}}, \vartheta_{\text{max}}]$. We start with an initial uniform partition into $\bar{\mu}$ subintervals with endpoints

$$\vartheta_{\mu} = \vartheta_{\text{min}} + \frac{\mu}{\bar{\mu}} (\vartheta_{\text{max}} - \vartheta_{\text{min}}), \quad \mu = 0, \ldots, \bar{\mu}, \quad (23)$$

and we apply the univariate passivity check by computing the SHH eigenvalues at each $\vartheta_{\mu}$, as discussed in Section V-A. An adaptive refinement process is then constructed, based on the following observations. Considering a single subinterval $\Theta_{\mu} = [\vartheta_{\mu}, \vartheta_{\mu+1}],$

- if $\psi(\vartheta) = 0$, $\forall \vartheta \in \Theta_{\mu}$, the model is not passive in this subinterval due to the presence of imaginary SHH eigenvalues. For each $\vartheta$, such eigenvalues delimit at least one frequency band $\Omega_i(\vartheta)$ where at least one singular value of the model $H(j\omega; \vartheta)$ exceeds one, with a local maximum $\tilde{\sigma}_i(\vartheta)$ occurring at frequency $\tilde{\omega}_i(\vartheta)$.
- if $\psi(\vartheta) > 0$, $\forall \vartheta \in \Theta_{\mu}$, the model is uniformly passive in $\Theta_{\mu}$.

Assuming now that only information at the edges $\{\vartheta_{\mu}, \vartheta_{\mu+1}\}$ of $\Theta_{\mu}$ is available, we have the following subcases (Fig. 2):

- if $\psi(\vartheta_{\mu}) = \psi(\vartheta_{\mu+1}) = 0$ and $\nu(\vartheta_{\mu}) = \nu(\vartheta_{\mu+1})$, we can infer that there exist at least one region $\Gamma_{\mu,i}$ of the frequency-parameter plane $\Gamma \subseteq [0, +\infty) \times \Theta_{\mu}$ where the model is uniformly non-passive (see Figure 2a). For each region $\Gamma_{\mu,i}$, we know the local singular value maxima $\tilde{\sigma}_i(\vartheta_{\mu}), \tilde{\sigma}_i(\vartheta_{\mu+1})$, and the corresponding localization frequencies $\tilde{\omega}_i(\vartheta_{\mu}), \tilde{\omega}_i(\vartheta_{\mu+1})$. In addition, we
know all the edges of $\Gamma_{\mu,i}$ for $\theta_\mu$ and $\theta_{\mu+1}$ from the sets $\chi(\theta_{\mu}), \chi(\theta_{\mu+1})$. Since local passivity violations have been identified, there is no need to refine subinterval $\Theta_{\mu}$.

- if $\psi(\theta_{\mu}) = \psi(\theta_{\mu+1}) = 0$ and $\nu(\theta_{\mu}) \neq \nu(\theta_{\mu+1})$, the number of frequency bands possibly changes when sweeping $\vartheta \in \Theta_{\mu}$ (see Figure 2b). In order to obtain a precise characterization, we refine $\Theta_{\mu}$ by adding the new point $\theta_{\mu+1/2} = 1/2(\theta_{\mu} + \theta_{\mu+1})$. (24)

- if $\psi(\theta_{\mu}) = 0$ and $\psi(\theta_{\mu+1}) > 0$, or conversely $\psi(\theta_{\mu}) > 0$ and $\psi(\theta_{\mu+1}) = 0$, we have a transition from a passive to a non-passive model while sweeping $\vartheta \in \Theta_{\mu}$ (see Figure 2c). Therefore, we need to refine $\Theta_{\mu}$ through (24) in order to track the particular $\vartheta_s$ where the onset of a passivity violation occurs.

- if both $\psi(\theta_{\mu}) > 0$ and $\psi(\theta_{\mu+1}) > 0$, we can have two subcases:
  - $\psi(\vartheta) > 0$ throughout $\Theta_{\mu}$: model is uniformly passive in $\Theta_{\mu}$ and no refinement is necessary (Figure 2d);
  - $\psi(\vartheta)$ vanishes in some subinterval $\Theta_s \subseteq \Theta_{\mu}$, denoting a passivity violation that is not visible from the edges of $\Theta_{\mu}$ (Figure 2e). We need to refine $\Theta_{\mu}$.

These two cases are here discriminated by computing $\psi(\theta_{\mu+1/2})$, and by constructing the first-order interpolation error

$$\varepsilon_{\mu+1/2} = \left| \psi(\theta_{\mu+1/2}) - 1/2[\psi(\theta_{\mu}) + \psi(\theta_{\mu+1})] \right|. \tag{25}$$

Refinement is applied if

$$\varepsilon_{\mu+1/2} > \gamma \left| \psi(\theta_{\mu+1/2}) \right|, \tag{26}$$

equivalently, when the estimate of the midpoint through linear interpolation is not sufficient to infer that $\psi(\vartheta)$ is uniformly positive within $\Theta_{\mu}$. The parameter $\gamma$ can be used to tune the selectivity of the refinement test, here we use the conservative value $\gamma = 0.2$.

We perform a total number $M$ of refinement passes. At each pass, the new points $\theta_{\mu+1/2}$ are added to the previous subset $\{\theta_{\mu}, \mu = 0, \ldots, m\}$, which is resorted in ascending order and reindexed, by suitably redefining $\bar{\mu}$. Algorithm 1 summarizes the proposed multivariate passivity check in pseudocode form (step 10 embeds the determination of new parameter samples $\theta_m^m$ at each refinement pass $m$, according to the above adaptive process). Throughout this work, we use $M = 10$.

A final remark is in order about the initial sampling (23). Here, we need to make sure that we do not miss important information due to an excessively coarse sampling. On the other hand, we do not want to spend unnecessary computations if the number of samples is too large. The number of initial subbands $\bar{\mu}$ depends in fact on the model variations throughout the parameter space, which in turn is directly related to the type and number of adopted basis functions $\xi_\ell(\theta)$ for model parameterization in (5). In this work, we assume entire-domain smooth basis functions (polynomials, orthogonal polynomials, or trigonometric polynomials). Due to the smooth parameterization, it is sufficient to consider the heuristic rule $\bar{\mu} = \kappa \bar{\ell}$, with $\kappa > 1$. For all documented examples, we set $\kappa = 4$.

### Algorithm 1: Passivity check of parameterized models

**Require:** frequency basis $\varphi_n$ for $n = 0, \ldots, \bar{n}$;

**Require:** parameter basis $\xi_\ell$ for $\ell = 1, \ldots, \bar{\ell}$;

**Require:** model coefficients $R_{n,\ell}, r_{n,\ell}$ in (2) and (5);

**Require:** control parameters $\vartheta_{\min}, \vartheta_{\max}, \gamma, \kappa, M$;

1. set $m = 0$ and number of initial samples $\mu_0 = \kappa \bar{\ell}$;
2. set initial samples $S_0 = \{\theta^m_0, \mu = 0, \ldots, \bar{\mu}_0\}$ as in (23);
3. repeat
   4. for $\mu = 1, \ldots, \bar{\mu}_m$ do
      5. construct SHH pencil $(\theta^m_\mu, \mathbf{K})$;
      6. find imaginary SHH eigenvalues $\omega_i(\theta^m_\mu)$;
      7. extract local singular value maxima $(\tilde{s}(\omega^m_{\mu,i}, \sigma^m_{\mu,i});$
      8. end for
   9. $m \leftarrow m + 1$
10. determine new samples $S_m = \{\theta^m_\mu, \mu = 1, \ldots, \bar{\mu}_m\}$;
11. until $S_m = \emptyset$ or $m = M$
12. return passivity violations $\mathcal{V} = \cup_m \{(\omega^m_{\mu,i}, \sigma^m_{\mu,i}); \forall \mu, i\}$.

Figure 3 illustrates the proposed passivity check on a practical example (discussed in more detail in Section VII-A). Two different models of a PCB interconnect link are processed by the proposed passivity check algorithm. Figure 3(a) plots the function $\psi(\vartheta)$ for the two models. The model no. 1 is not passive for any value of $\vartheta$, hence $\psi(\theta_0) = 0$ for all computed $\mu$ (the red dots). Correspondingly, Figure 3(b) shows that for all $\theta_\mu$ samples, there is at least one imaginary SHH eigenvalue (yellow dots) that defines a frequency band where at least one singular value of the model is larger than one (red lines). Local singular value maxima are close to DC (black squares). Note that, since the number of detected imaginary SHH eigenvalues is constant for all $\theta_\mu$, no adaptive refinement is necessary. The blue curve in Figure 3(a) shows instead $\psi(\theta)$ for model no. 2, which is locally not passive only in a restricted range of $\vartheta \in \Theta_s \approx [437, 454] \mu m$. Correspondingly, imaginary SHH eigenvalues are only found in this range: see Figure 3(c), where adaptive refinement is clearly visible near the transitions between $\psi(\theta) = 0$ and $\psi(\theta) > 0$, which define the endpoints of $\Theta_s$. Finally, Figure 3(d) shows the frequency-dependent singular values of model no. 1 for $\vartheta = 400 \mu m$, where the frequency of the unique imaginary SHH eigenvalue is highlighted by a yellow dot.

### VI. ENFORCING PASSIVITY OF PARAMETERIZED MODELS

Based on the multivariate passivity characterization discussed in Section V-B, we are now ready to formulate our proposed passivity enforcement algorithm. A perturbed model based on (9) is defined as

$$\hat{\mathbf{H}}(s; \vartheta) = \mathbf{H}(s; \vartheta) + \Delta \mathbf{H}(s; \vartheta) \tag{27}$$

with

$$\Delta \mathbf{H}(s; \vartheta) = \frac{\Delta \mathbf{N}(s, \vartheta)}{\mathbf{D}(s, \vartheta)} = \sum_{n=0}^{\bar{n}} \sum_{\ell=1}^{\bar{\ell}} \Delta R_{n,\ell}^n \xi_\ell(\theta) \varphi_n(s) \frac{1}{\sum_{n=0}^{\bar{n}} \sum_{\ell=1}^{\bar{\ell}} r_{n,\ell} \xi_\ell(\theta) \varphi_n(s)} \tag{28}$$

The coefficient perturbations $\Delta R_{n,\ell}$ will be computed such that the singular values of the perturbed model that are larger...
than one will be displaced below the unit threshold. Such coefficients will be collected in a global vector of decision variables

\[ \mathbf{x} = [\mathbf{x}_{i,j}^T, \cdots, \mathbf{x}_{P,p}^T]^T, \]  

where

\[ \mathbf{x}_{i,j} = [(\Delta R_{0,1})_{i,j}, \cdots, (\Delta R_{n,\ell})_{i,j}, \cdots, (\Delta R_{n,\bar{\ell}})_{i,j}]^T \]  

(30)

collects the \((i, j)\) entries of all matrices \(\Delta R_{n,\ell}\) with a suitable ordering. We have \(\mathbf{x}_{i,j} \in \mathbb{R}^{(n+1)\ell}\) and \(\mathbf{x} \in \mathbb{R}^Q\), with \(Q = P^2(n+1)\ell\).

A. Building Algebraic Passivity Constraints

Let us consider a single local singular value maximum \(\bar{\sigma}_{\mu,i} = \bar{\sigma}_i(\theta_\mu) > 1\) occurring at frequency \(\bar{\omega}_{\mu,i} = \omega_i(\theta_\mu)\), as resulting from the passivity check of Section V-B. Model perturbation leads to a perturbation of this singular value, which under first-order approximation reads [49]

\[ \bar{\sigma}_{\mu,i} \approx \bar{\sigma}_{\mu,i} + \Re \left\{ \mathbf{u}_{\mu,i}^H \Delta \mathbf{H}(j\bar{\omega}_{\mu,i}; \theta_\mu) \mathbf{v}_{\mu,i} \right\}, \]  

(31)

where \(\mathbf{u}_{\mu,i}, \mathbf{v}_{\mu,i}\) are the left and right singular vectors of \(\mathbf{H}(j\bar{\omega}_{\mu,i}; \theta_\mu)\) associated to \(\bar{\sigma}_{\mu,i}\). Forcing this singular value to be less than one leads to the inequality constraint

\[ \Re \left\{ \mathbf{u}_{\mu,i}^H \Delta \mathbf{H}(j\bar{\omega}_{\mu,i}; \theta_\mu) \mathbf{v}_{\mu,i} \right\} \leq 1 - \bar{\sigma}_{\mu,i}. \]  

(32)

Defining now

\[ \mathbf{a}_{\mu,i} = [(a_{\mu,i,0,1}, \cdots, a_{\mu,i,n,\ell}, \cdots, a_{\mu,i,n,\bar{\ell}})]^T \]  

(33)

with the same ordering as in (30), where

\[ a_{\mu,i,n,\ell} = \frac{\xi_\ell(\theta_\mu) \varphi_n(j\bar{\omega}_{\mu,i})}{D(j\bar{\omega}_{\mu,i}, \theta_\mu)}, \]  

(34)

we can cast (32) in algebraic form

\[ \mathbf{p}_{\mu,i}^T \mathbf{x} \leq 1 - \bar{\sigma}_{\mu,i}, \]  

(35)

where

\[ \mathbf{p}_{\mu,i}^T = \Re \left\{ (\mathbf{v}_{\mu,i}^T \otimes \mathbf{u}_{\mu,i}^H) \otimes \mathbf{a}_{\mu,i} \right\}. \]  

(36)

The constraint (32) operates on an individual singular value maximum \(\bar{\sigma}_{\mu,i}\), attempting to reduce its value to be less than one (since based on a first-order singular value perturbation, this enforcement is not exact but only approximate). If multiple singular value maxima are perturbed concurrently, it is sufficient to add a constraint (32) for each \(\mu, i\).

B. Preserving Model Accuracy

Enforcing constraint (32) does not guarantee that the perturbed model remains accurate. Hence the need of constructing a cost function that casts in algebraic form a suitable norm of the model perturbation. We define such cost function as

\[ \mathcal{E}^2 = \sum_{i,j=1}^P \mathcal{E}_{i,j}^2, \]  

(37)

where

\[ \mathcal{E}_{i,j}^2 = \sum_{k=1}^k \sum_{m=1}^m w_{i,j;k,m}^2 |\Delta H_{i,j}(j2\pi f_k, \vartheta_m)|^2 \]  

(38)

denotes the \((i, j)\)-th entry of the squared model perturbation at frequency \(f_k\) and parameter \(\vartheta_m\), weighted by the possibly frequency-, parameter-, and entry-dependent weight \(w_{i,j;k,m}\). Defining now

\[ b_{k,m;n,\ell} = \frac{\xi_\ell(\vartheta_m) \varphi_n(j2\pi f_k)}{D(j2\pi f_k, \vartheta_m)}, \]  

(39)

and

\[ \mathbf{b}_{k,m}^T = [b_{k,m,0,1}, \cdots, b_{k,m,n,\ell}, \cdots, b_{k,m,n,\bar{\ell}}] \]  

(40)
we can cast the elementwise cost function (38) as
\[ E_{i,j}^2 = \| F_{i,j} x_{i,j} \|_2^2, \]  
(41)

where
\[
F_{i,j} = \begin{bmatrix} \operatorname{Re}(\tilde{F}_{i,j}) \\ \operatorname{Im}(\tilde{F}_{i,j}) \end{bmatrix}, \quad \tilde{F}_{i,j} = \begin{bmatrix} w_{i,j;1,1} b^T_{1,1} \\ \vdots \\ w_{i,j;k,m} b^T_{k,m} \\ \vdots \\ w_{i,j;k,m} b^T_{k,m} \end{bmatrix}
\]
(42)

Note that \( F_{i,j} \in \mathbb{R}^{2km \times (n+1)\ell} \) collects as many rows as available frequency and parameter samples. The row size can thus be large, since usually \( 2km \gg (n+1)\ell \). An equivalent compressed form of (41) reads
\[ E_{i,j}^2 = \| \Psi_{i,j} x_{i,j} \|_2^2, \]  
(43)

where \( \Psi_{i,j} \in \mathbb{R}^{(n+1)\ell \times (n+1)\ell} \) is obtained through an “economy-size” QR factorization of \( F_{i,j} = Q_{i,j} \Psi_{i,j} \), where \( Q_{i,j} Q_{i,j}^T = I \). Finally (37) can be cast as
\[ E^2 = \| \Psi x \|_2^2, \]  
(44)

where
\[ \Psi = \text{blkdiag}(\Psi_{i,j})_{1 \leq i,j \leq 1} \]  
(45)

C. Iterative Passivity Enforcement

A first-order singular value perturbation with minimum induced model error is achieved by minimizing (44) while enforcing (35), resulting in the following constrained minimization problem
\[ \min_{x} \| \Psi x \|_2^2 \quad \text{subject to} \quad P_{\mu,i}^T x \leq 1 - \sigma_{\mu,i}, \quad \forall \mu, i. \]  
(46)

This problem is convex thus straightforward to solve, e.g., through a Primal-Dual Interior Point method [50]. Since based on a first-order approximation, the above optimization may need to be iterated until all passivity violations are removed. Algorithm 2 illustrates the resulting passivity enforcement scheme in pseudocode form. We see that this algorithm is a straightforward extension to the multivariate case of the singular value perturbation scheme [26], which is only applicable to univariate models. The key for the proposed multivariate extension is the availability of a reliable process for detecting and extracting the parameter-dependent passivity violations of the model, as discussed in Section V.

VII. EXAMPLES

The performance of proposed passivity enforcement scheme is now demonstrated on three examples. A laptop with Intel Core i7 CPU running at 2.6 GHz with 16 GB RAM was used in all numerical simulations.

Algorithm 2 Passivity enforcement of parameterized models

\begin{algorithm}
  \begin{algorithmic}[1]
    \Require frequency basis \( \varphi_n \) for \( n = 0, \ldots, \bar{n} \);
    \Require parameter basis \( \xi_{\ell} \) for \( \ell = 1, \ldots, \bar{\ell} \);
    \Require model coefficients \( R_{n,\ell}, r_{n,\ell} \) in (2) and (5);
    \Require control parameters \( \theta_{\min}, \theta_{\max}, \gamma, \kappa, M \);
    \State find passivity violations \( \mathcal{V} \) via Algorithm 1;
    \While {\( \mathcal{V} \neq \emptyset \)}
      \State build constraint (35) for each element in \( \mathcal{V} \);
      \State compute matrix \( \Psi \) in (45);
      \State solve convex optimization problem (46);
      \State update model coefficients \( R_{n,\ell} \leftarrow R_{n,\ell} + \Delta R_{n,\ell} \);
    \EndWhile
    \State find passivity violations \( \mathcal{V} \) via Algorithm 1;
    \Return passive model \( \hat{H}(s; \theta) \).
  \end{algorithmic}
\end{algorithm}

A. A Printed Circuit Board Interconnect

The first example we consider is a high-speed signal link (see [51] for a detailed description) routed on the inner layers of two Printed Circuit Boards hinged by a connector. Vertical interconnection at the feeding ports and at the connector ports is provided by four through vias. The 2 \( \times \) 2 scattering responses of the link are parameterized by the via antipad radius \( \vartheta \neq r \) within the range \( \Theta = [400, 600] \mu m \). The raw frequency-domain scattering responses (Courtesy of Prof. Christian Schuster and Dr. Jan Preibisch, Technische Universität Hamburg-Harburg, Hamburg, Germany) are obtained through a combination of a full-wave field solver (for the connector), lossy transmission-line models for the stripline segments, and a field model for the vias based on [52]. A total of \( \bar{k} = 500 \) frequency samples up to 10 GHz, combined with \( \bar{m} = 9 \) parameter samples were available for model identification.

The parameterized GSK iteration of Section II was applied with \( \bar{n} = 44 \) basis poles, and using orthogonal (Chebyshev) polynomials \( \xi_{\ell}(\vartheta) \) to represent parameter variations through (5). The number of basis functions was determined by trial and error as \( \ell = 3 \), corresponding to quadratic polynomials. Only one half of the parameter samples (indices \( m = 1, 3, 5, 7, 9 \)) were used to fit the model, leaving the even-numbered \( m = 2, 4, 6, 8 \) for a-posteriori model validation purposes. The worst-case absolute RMS error among all scattering matrix elements at both fitting and validation points resulted 8.788 \( \times \) \( 10^{-4} \), whereas the relative error was 3.367 \( \times \) \( 10^{-3} \). This model corresponds to the model at Iteration 1 depicted in Fig. 3. We see from panels (a), (b), and (d) that this model is not passive, due to localized low-frequency passivity violations throughout the parameter range.

The proposed passivity enforcement algorithm required three iterations and a runtime of 20 seconds. The final model, as evident from Fig. 4, is uniformly passive since \( \psi(\vartheta) > 0 \) for all \( \vartheta \in \Theta \). Figure 5 shows a comparison between the model responses and the raw scattering samples for all fitting and validation points. We see that the accuracy is excellent, as confirmed by the worst-case RMS errors 8.792 \( \times \) \( 10^{-4} \) (absolute) and 3.368 \( \times \) \( 10^{-3} \) (relative). These errors are only marginally worse than the errors of the original non-passive
model, thanks to the adopted cost function (44) that is minimized through passivity enforcement. This level of accuracy can be achieved only when the original passivity violations are very small (the worst-case passivity violation corresponded to a maximum singular value $\sigma_{\text{max}} = 1.000346$, which was larger than one by a very small amount). Finally, we depict the resulting implicitly-parameterized macromodel poles in Fig. 6, computed by instantiating the model at the original parameter samples $\vartheta_m$.

Fig. 4. Samples of $\psi(\vartheta)$ for the PCB link macromodel after passivity enforcement.

Fig. 5. Scattering responses $S_{11}(j\omega; \vartheta)$ of the PCB link macromodel after passivity enforcement (dashed red lines), compared to the raw data used for model identification (solid blue lines). Both fitting and validation points are displayed; the arrows denote the increasing direction for via antipad radius $\vartheta$.

Fig. 6. Parameterized poles (zoomed view) of the PCB link macromodel, computed and superimposed for all original parameter samples $\vartheta_m$.

Fig. 7. Localization of non-passive bands (red lines) through imaginary SHH eigenvalues (yellow dots) of original non-passive inductor model. Top and bottom panels zoom on two different regions of the frequency-parameter plane.

Fig. 8. Normalized SHH spectral distance from imaginary axis for the integrated inductor example, before (Iteration 1) and after (Iteration 6) passivity enforcement.

B. An integrated inductor

We consider here a 2-port integrated inductor (courtesy of Prof. Madhavan Swaminathan, Georgia Institute of Technology, Atlanta, USA). The inductor has a square outline with 1.5 turns routed on two different layers of the substrate. The scattering responses were obtained through a full-wave field solver for $\bar{m} = 11$ different values of the sidelength $L = \vartheta$ ranging from 1.02 to 1.52 mm, with $\bar{k} = 477$ frequency samples up to 12 GHz. Model identification via GSK iteration was performed using only odd-indexed parameter values, leaving the even-numbered responses for model validation purposes, with $\bar{n} = 8$ poles and $\bar{l} = 4, 3$ Chebychev polynomial basis functions for numerator and denominator, respectively.

The initial model was characterized by various passivity violations, illustrated in Fig. 7. These violations required a total of 5 iterations (runtime 15 seconds) to be removed, resulting in a uniformly passive model throughout the parameter range. Removal of passivity violations is confirmed by Fig. 8, which depicts $\psi(\vartheta)$ before and after passivity enforcement. The RMS errors of original and passive models with respect to the original scattering responses are reported in Table I, showing that both models are very accurate. Figure 9 confirms this statement by comparing passive model responses to raw data for all fitting and validation points.
C. A filter

We consider now a double-folded microstrip filter (see [53] for a more detailed description), which can be tuned by changing the length \( \bar{\vartheta} = \frac{L}{k} \) of a microstrip stub within the range 2.08–2.28 mm. Scattering responses for \( m = 21 \) linearly-spaced stub length values were computed through a field solver at \( k = 300 \) samples spanning the frequency band [5, 20] GHz. Model identification based on odd-indexed parameter samples required \( \bar{m} = 10 \) poles and \( \bar{l} = 3 \) (quadratic) polynomial basis functions to capture parameter variations.

The resulting model was not passive, as Fig. 10 shows by depicting \( \psi(\bar{\vartheta}) = 0 \) at the first passivity iteration loop.

The localized passivity violations of the initial model are depicted in the top panel of Fig. 11. Passivity enforcement required 5 iterations (runtime 31 seconds) to achieve uniform passivity, as confirmed by \( \psi(\bar{\vartheta}) > 0 \) in Fig. 10. The passivity violations at the second iteration are depicted in the bottom panel of Fig. 11, where we can note that the extent of such violations both in the frequency and parameter directions is reduced. Note also that the presence of multiple singular values exceeding one (Fig. 11, Iteration 1, localized at 22–24 GHz and approximately \( \bar{\vartheta} \in [2.08, 2.13] \) mm), does not pose particular problems, since multiple independent constraints can be enforced while solving (46).

Finally, Fig. 12 compares the model responses to the original scattering responses at the validation points (even-indexed parameter samples). Also for this example the accuracy is excellent, with a worst-case RMS error amount all parameter, frequency samples and scattering matrix elements equal to \( 3.16 \times 10^{-3} \) (absolute) and \( 4.64 \times 10^{-3} \) (relative).
D. A partially coupled interconnect

The last example we discuss is a more complex 6-port structure. Six circular wires (radius $r_w = 0.5$ mm, dielectric insulation radius $r_d = 0.8$ mm with relative permittivity $\varepsilon_R = 4.2$) are arranged in a $2 \times 3$ matrix, with center-to-center separation of adjacent wires $D = 1.61$ mm. The wires form a set of 3 differential pairs of total length $L = 10$ cm. The three pairs are coupled over a length $\vartheta = L_c$, which is a free parameter in the range $\vartheta \in \Theta = [20, 40]$ mm. For the remaining length $L - L_c$ two pairs are coupled and the third pair is uncoupled. The structure is modeled as a network of uniform transmission lines, whose RLGC parameters are extracted using a 2D field solver and used to evaluate the $6 \times 6$ differential scattering matrix of the three differential pairs. Differential ports 1–3 are on the fully coupled near-end side, whereas ports 4–5 (coupled) and 6 (uncoupled) are on the far-end side.

A parameterized macromodel was identified starting from $m = 11$ linearly spaced parameter samples times $k = 500$ frequency samples up to 5 GHz, using $n = 30$ poles and $l = 5$ (degree-4) Chebychev polynomials for numerator and denominator submodels. The size of the SHH pencil is 372 for this case, which is then considered as a medium-scale problem.

Figure 13 demonstrates that the proposed algorithm is able to track and remove complex and overlapping passivity violation regions, characterized by multiple singular values larger than one; note that the number of singular values is equal to the number of ports $P$, which is thus the upper bound of the number of pointwise passivity violations in the $(\omega, \vartheta)$ plane. As Fig. 14 shows, after 10 iterations of the passivity enforcement loop the condition $\psi(\vartheta) > 0$ was achieved $\forall \vartheta = L_c$ in its range, confirming uniform passivity of the final model. The accuracy of the passive model is demonstrated in Fig. 15 on one of the sensitive far crosstalk responses between two partially-coupled differential pairs (with a worst-case RMS error $2.9 \times 10^{-3}$ between passive model responses and raw data).

VIII. CONCLUSION

This paper presented an efficient and robust algorithm for checking and enforcing the passivity of behavioral macromodels of LTI systems, whose scattering matrix depends both on frequency and on one additional external parameter. Thanks to a specialized adaptive sampling process in the parameter space, based on the spectral properties of some
Skew-Hamiltonian/Hamiltonian matrix pencil associated to the model, we are able to detect and localize all frequency- and parameter-dependent passivity violations of the model. An iterative constrained optimization loop is able to remove such violations, while retaining model accuracy. To be best of Author’s knowledge, this is the first documented algorithm that is able to enforce the uniform passivity of a parameterized macromodel.

Our current implementation of the parameterized modeling flow is applicable to small- and medium-scale problems, since the computational cost grows very fast when increasing dynamic order (number of poles), number of parameters with the associated polynomial degree of numerator/denominator submodels, and amount of raw data (related quadratically to the number of ports). The most critical step is actually the associated polynomial degree of numerator/denominator submodels, and amount of raw data (related quadratically to the number of ports). The most critical step is actually model identification via SK iteration, which requires solving a large-size and dense linear system at each iteration. The passivity enforcement scheme that is presented in this work offers instead better scalability properties, since it is based on an iterative adaptive processing of standard univariate models.

Future research will be devoted to the extension of proposed approach to higher dimensions in the parameter space, with special emphasis on avoiding issues due to the curse of dimensionality. This task will require a suitable generalization of the proposed adaptive sampling strategy to higher dimensions, as well as a reformulation of the SK iteration that will ensure a favorable scalability.

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