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## Incoming windblown sand drift to civil infrastructures: a probabilistic evaluation

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#### Abstract

The accurate prediction of windblown sand drift events approaching human infrastructures and activities is fundamental in arid lands. In both scientific literature and technical practice sand drift estimation is carried out in mean terms. Typically, sand drift net direction and intensity are assessed by means of the resultant drift potential. However, windblown sand suffers a number of epistemic and aleatory uncertainties, related to both the wind and the sand fields. The windblown sand drift estimation in probabilistic terms is useful in the infrastructure design perspective and allows to obtain characteristic values of windblown sand transport. In this study windblown sand is considered as an environmental action in analogy to wind action. Several uncertainties involved in the phenomenon are considered: threshold shear velocity and 10-minute average wind velocity are assumed as random variables. Monte Carlo approach is adopted within a bootstrapping technique in order to assess sand drift in probabilistic terms. The proposed approach is applied to five sites in the Arabian Peninsula. Directional statistics of the sand drift are given for each site.

Keywords: windblown sand, drift potential, uncertainty quantification, probabilistic approach, Monte Carlo

#### Nomenclature

DP	Drift Potential
HW	Hybrid Weibull
MC	Monte Carlo
RDD	Resultant Drift Direction
RDP	Resultant Drift Potential
SD-WA	Sand Deterministic - Wind Averaged
SWP	Sand Wind Probabilistic
D	drift potential
F	probability distribution function
$F_0$	wind calm rate
Ν	number of occurrences
Q	sand transport rate
R	resultant drift potential
Т	reference time
$T_r$	recording time
U	wind velocity
$U_{10}$	10-min averaged wind speed

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<i>C.O.V.</i>	coefficient of variation
d	sand grain diameter
$d_r$	sand grain reference diameter
f	probability density function
g	gravitational acceleration
k	Weibull shape parameter
р	percentile
sk	skewness
$u_*$	shear velocity
$u_{*t}$	threshold shear velocity
$z_0$	roughness length
$\Delta t$	sampling interval
$\Delta \theta$	sector width
$\theta$	wind direction
λ	Weibull scale parameter
$\mu$	mean value
$ ho_a$	air density
$ ho_b$	packed bulk sand density
$\sigma$	standard deviation
#	cardinality
0	calm wind

#### 1 1. Introduction

Windblown sand is of interest for several engineering fields in arid environments (e.g. Middleton and Sternberg, 2 2013; Stipho, 1992), from environmental to civil engineering. In particular, windblown sand interacts with a number 3 of civil structures and infrastructures, such as roads (e.g. Redding and Lord, 1981; Dong et al., 2004), railways (e.g. Zhang et al., 2007, 2010; Cheng and Xue, 2014), industrial facilities and pipelines (e.g. Alghamdi and Al-Kahtani, 5 2005), farms (e.g. Wang et al., 2010), town an buildings (e.g. Rizvi, 1989; Bofah and Al-Hinai, 1986). Windblown 6 sand transport results from soil erosion and involves sedimentation around built obstacles. In particular, windblown 7 sand effects on civil structures comprehend, but are not limited to: wind erosion and foundation scouring, moving sand 8 dunes encroaching infrastructures, sand accumulation around structures and infrastructures. Due to the nature of these 9 effects, they can lead to several incremental costs in infrastructure management, e.g. loss of capacity and increased 10 maintenance costs (Zakeri, 2012), but also to disastrous events, such as train derailment (Cheng et al., 2015). The 11 design of such infrastructures requires the accurate estimation of the amount of incoming windblown sand that attacks 12 the structure. It significantly vary in space and time. Indeed, on the one hand, line-like infrastructures cross different 13 regions with a wide variety of geomorphological characteristics. On the other hand, infrastructure design must ensure 14 the service life prescribed by standards. Hence, a probabilistic approach to design is necessary to take into account 15 the inborn variability of the phenomenon. 16

The amount of incoming windblown sand is defined as the mass per unit time and per unit length, and usually 17 called *incoming sand drift*. Phenomenologically, windblown sand is a multi-physics phenomenon which includes wind 18 and sand subfields. Hence, sand drift depends on both the wind velocity and the sand characteristics. The modelling 19 framework to sand drift evaluation has been first introduced by Fryberger and Dean (1979). Their seminal work 20 still grounds the current scientific and technical literature in several application fields, such as fundamental research 21 (e.g. Al-Awadhi and Al-Awadhi, 2009; Barchyn and Hugenholtz, 2011), geomorphology (e.g. del Valle et al., 2008; 22 Bogle et al., 2015; Kilibarda and Kilibarda, 2016; Yang et al., 2016), paleo sedimentology (e.g. Yang et al., 2014), 23 climatology (e.g. Bogle et al., 2015), coastal management (e.g. Riksen et al., 2016), civil engineering (e.g. Dong 24 et al., 2004; Zhang et al., 2010; Cheng et al., 2015). In the Fryberger and Dean (1979) framework, the so-called Drift 25 Potential (DP) is defined for each wind direction, while the Resultant Drift Potential (RDP) and the Resultant Drift 26 Direction (RDD) stand for the magnitude and direction of the vector sum of DP over the directions, respectively. 27 These quantities are called "potential" because they provide a measure of sand-moving capacity of the wind blowing 28

over an ideal sand bed, neglecting the local covering of the ground surface (Pye and Tsoar, 2009). Fryberger and 29 Dean (1979) obtain DP per reference time (usually 1 year) by cumulating the sand transport rate Q over the wind 30 speed recording time, and rescaling it on the reference time. In turn, Q results from the vertical integration of the 31 horizontal windblown sand flux. Several semi-empirical models to predict Q have been proposed so far, reviewed e.g. 32 in Dong et al. (2003); Kok et al. (2012); Sherman and Li (2012). Among them, modified Bagnold type models are 33 the most widely adopted in literature (see for instance the field studies by Fryberger and Dean, 1979; Al-Awadhi and 34 Al-Awadhi, 2009; Barchyn and Hugenholtz, 2011; Sherman and Li, 2012; Sherman et al., 2013; Yang et al., 2014; Liu 35 et al., 2015). In particular, the model proposed by Lettau and Lettau (1978) is the most adopted one. They all relate Q36 to the wind shear velocity  $u_*$  and the threshold shear velocity  $u_{*t}$ , that is the shear velocity above which sand transport 37 occurs. Usually, such a threshold is assessed as a function of the sand grain diameter d by means of semi-empirical 38 expressions (e.g. Bagnold, 1941; Iversen and White, 1982; Shao and Lu, 2000; McKenna, 2003). According to the 39 Authors, it is worth pointing out that the current approach within the Fryberger and Dean (1979) framework is: 40

- deterministic with respect to the sand subfield. Indeed, the expressions of the threshold shear velocity  $u_{*t}$  used so far are purely deterministic;
- time-averaged with respect to the wind subfield. The wind speed inborn variability is accounted for, but only the mean value of DP is retained because the rescaling on the reference time is tantamount to averaging.

<sup>45</sup> Let us call such approach as Sand Deterministic - Wind Averaged (SD-WA).

Despite SD-WA approach is generalized in practice, windblown sand phenomenon is affected by several sources 46 of uncertainty. They can be generally classified in *aleatory* and *epistemic* uncertainties (Zio and Pedroni, 2013). Let 47 us introduce a complementary categorization referring to the wind and sand subfields introduced above. Epistemic uncertainties are associated with the lack of knowledge about the properties and conditions of the phenomena to be 49 modeled. They can be further ascribed to model, parameter and measurement uncertainties. Wind-field epistemic 50 uncertainties are generally well quantified, because of its long-standing modelling, while sand-field ones have been 51 only recently highlighted with respect to threshold shear velocity (e.g. Barchyn and Hugenholtz, 2011; Raffaele et al., 52 2016) and sand transport rate (e.g. Barchyn et al., 2014). Aleatory uncertainties refer to inherent randomness of natural 53 phenomena. Let us introduce a further categorization referring to the wind and sand subfields introduced above. Wind-54 related aleatory uncertainties affect the velocity and other environment variables. Sand-related aleatory uncertainties 55 take place at both the microscopic scale, i.e. grain irregular shape, grain size distribution, grain relative position on the sand bed (e.g. Nickling, 1988; Duan et al., 2013; Edwards and Namikas, 2015), and the macroscopic scale, i.e. 57 soil vegetation covering, soil sediment availability, soil moisture and soil crusting (see e.g. McKenna Neuman and 58 Nickling, 1989; Lancaster and Baas, 1998; Shao, 2008; Hoonhout and de Vries, 2016). The statistical description of 59 wind speed is long-standing and well established, as reviewed e.g. by Carta et al. (2009). Conversely only recently 60 the Authors proposed the statistical description of threshold shear velocity (e.g. Raffaele et al., 2016). The cited paper 61 substantially contributes to the background of the present study. It includes a comprehensive review on the uncertain-62 ties that affects both experimental measurements and modelling of  $u_{*t}$ . In the light of this, a statistical modelling is 63 developed, based on advanced copula-based quantile regression. Joint probability density functions of the sand grain 64 diameter and  $u_{*t}$  are derived, as well as the conditional probability density functions of the threshold shear velocity 65

<sup>66</sup> for given values of the diameter.

Both the engineering design needs and the shortcomings of the SD-WA approach pave the way for the proba-67 bilistic description of the incoming sand drift. According to the Authors, it can be regarded as equivalent to other 68 environmental actions, in analogy to wind action. Hence, let us briefly outline in the following to which extent the 69 incoming wind speed U is analogous to the sand transport rate Q and to the drift potential DP. First, in wind engi-70 neering the wind speed is defined in probabilistic terms due to the uncertainty related to inborn wind variability only. The probabilistic representation of sand transport rate is recommended a fortitiori and it is more difficult at the same 72 time, since it is affected by more uncertainties. The variability of both wind and sand features should be taken into 73 account. Second, most of the wind effects on structures, e.g. equivalent static loads or flutter, are related to extreme 74 75 values of the incoming wind speed. Conversely, windblown sand effects on civil structures are mainly induced by the cumulated values of current values of Q over time, that is DP. In this perspective, windblown sand effects and related 76 assessment recall wind-induced fatigue. In spite of this analogy, some differences remain. Only a few incoming wind 77 directions are considered in directional wind-induced fatigue assessment (see e.g. Repetto and Solari, 2004), i.e. the 78

<sup>79</sup> ones that induce the highest stresses on the cross section. Conversely, all incoming wind directions are taken into <sup>80</sup> account in assessing the windblown sand drift, since they all contribute in RDP definition.

Bearing the above analogy in mind, three main questions may rise to the Authors' mind: i. How does the uncertainty of both threshold shear velocity and mean wind velocity jointly propagate to RDP? ii. Does the probability distribution of RDP change significantly form a site to another in the same region? iii. Does the gap between characteristic and mean value of RDP make the approach of interest for engineering practice?

The present study aims at contributing in shedding some light on such issues. A general probabilistic approach 85 proposed and applied to real world sites in Arabian Peninsula. Each site is characterized by its actual wind field is 86 and sand granulometry. In particular, variability of both sand characteristics at microscopic scale (comprehensively 87 reflected by threshold shear velocity) and wind speed (i.e. wind direction and intensity) are considered. Other sources 88 of uncertainties reviewed above are not included because the lack of their statistical description. As a result, instead 89 of a single pair of values describing mean RDP magnitude and direction, their probability distributions are obtained. 90 Characteristic values are derived from them, and design values can be derived in turn towards a semi-probabilistic 91 approach. The paper develops accordingly to the above objectives through the following sections. In Section 2, the 92 proposed probabilistic approach is outlined. In Section 3, results referred to some chosen Sites in Arabian Peninsula 93 are shown, compared and discussed. In Section 4, conclusions and perspectives are outlined. 94

#### 95 2. Methods

In the following, the proposed probabilistic approach based on the general framework of Fryberger and Dean (1979) is outlined. First, the framework proposed by Fryberger and Dean (1979) is recalled. Then, the proposed probabilistic approach to assess sand drift is shown step-by-step.

<sup>99</sup> Fryberger and Dean (1979) define the directional drift potential and the resultant drift potential on the basis of the <sup>100</sup> model proposed by Lettau and Lettau (1978), where the sand transport rate  $Q_{\theta}$  in a given direction  $\theta$  is expressed as

being d the sand grain diameter,  $d_r = 0.25$  mm the reference sand grain diameter,  $\rho_a$  the air density, g the gravitational

<sup>102</sup> acceleration,  $u_{*t}$  the threshold shear velocity and  $u_{*,\theta}$  the shear velocity in the corresponding wind direction.

<sup>103</sup> The directional drift potential  $D_{\theta}$  (i.e. DP in Fryberger and Dean, 1979, notation) is rephrased as

$$D_{\theta} = \frac{1}{\rho_b} \frac{T}{T_r} \sum_{i=1}^{N_{\theta}} Q_{\theta,i} \Delta t = \frac{T}{T_r} \sum_{i=1}^{N_{\theta}} D_{\theta,\Delta t,i} \quad (\text{or } D_{\theta} = 0 \text{ if } N_{\theta} = 0), \qquad (2)$$

where  $\rho_b$  is the packed bulk sand density, T is the reference time and  $T_r$  is the recording time set as a multiple of T.

 $\Delta t$  is the sampling interval of the wind speed, not necessarily equal to the 10-minute averaging time, for the sake of

generality. The drift potential over the sampling interval  $D_{\theta,\Delta t} [m^3 m^{-1} \Delta t^{-1}]$  is estimated postulating  $Q_{\theta} [Kg m^{-1}s^{-1}]$ constant over  $\Delta t$ .

 $N_{\theta}$  follows as the number of occurrences in the reference time in which the wind will blow in the direction  $\theta$ , and it is expressed as

$$N_{\theta} = \frac{T}{T_r} \frac{T_{\theta}}{\Delta t},$$
 constrained by  $\sum_{\theta=1}^{2\pi/\Delta\theta} N_{\theta} + N_0 = N,$  (3)

where  $T_{\theta}$  is the time over which the wind blows in the direction  $\theta$ ,  $\Delta \theta$  is the sector width on which the wind is recorded,

 $N_0$  and N are the number of occurrences of calm wind and the number of total occurrences in the reference time T, respectively.

Finally, the resultant drift potential *R* (i.e. RDP in Fryberger and Dean, 1979, notation) can be easily obtained from the vector sum of  $D_{\theta}$ :

$$R = \sum_{\theta=1}^{2\pi/\Delta\theta} D_{\theta}.$$
(4)

In the following, resultant drift potential magnitude and direction are defined as |R| and  $\hat{R}$ , respectively.

It may be useful to highlight that Fryberger and Dean (1979) provide also an index of the directional variability of
 windblown sand drift, i.e. the ratio between the resultant drift potential magnitude and the sum of drift potential
 modulus:

$$R/D = \frac{|R|}{\sum_{\theta=1}^{2\pi/\Delta\theta} |D_{\theta}|}.$$
(5)

<sup>119</sup> In particular, the lower the ratio, the higher the directional variability.

In the proposed probabilistic approach the input quantities  $u_{*t}$  and  $u_{*,\theta}$  are random variables. Hence, the Fryberger

and Dean (1979) framework has to be adapted in order to deal with such random variables. Let us call such approach as Sand Wind Probabilistic (SWP). The steps followed in SWP approach are sketched in the flow chart in Figure 1

and described in the following.

123



Figure 1: Flow chart of the proposed SWP approach

The site characteristics are needed as input data, with respect to both sand subfield (mean sand diameter d) and 124 wind subfield (aerodynamic roughness  $z_0$  and time series of 10-minute averaged wind speed  $U_{10}(t)$ ). The input ran-125 dom variables  $u_{*t}$  and  $u_{*,\theta}$  are described from the probability density functions  $f(u_{*t})$  and  $f(u_{*,\theta})$ , respectively. To the 126 Authors' best knowledge, there are no experimental evidence or systematic studies in literature about a dependence 127 between  $u_{*,\theta}$  and  $u_{*t}$ . In this study, the directional shear velocity and the threshold shear velocity are considered in-128 dependent random variables. Indeed,  $u_{*t}$  depends entirely on the sand characteristics, while  $u_{*\theta}$  depends only on the 129 wind velocity for a given  $z_0$ .  $f(u_{*,\theta})$  is simply obtained by rescaling the probability density function  $f(U_{10,\theta})$ , being 130  $u_{*,\theta} = 0.41 U_{10,\theta} / ln(z/z_0)$ . Hence, Weibull-type  $f(u_{*,\theta})$  results. The conditional probability density functions  $f(u_{*t} | d)$ 131 are obtained in Raffaele et al. (2016). Interested readers can refer to the paper above for further details. Here, Figure 132 2 is limited to summarize the final finding of that study, i.e. the statistical description of the threshold shear velocity 133 versus mean sand diameter d by means of some percentiles and statistical moments: the mean value  $\mu(u_{*t})$  and  $1^{st}$ , 134  $5^{th}$ ,  $25^{th}$ ,  $75^{th}$ ,  $95^{th}$  and  $99^{th}$  percentiles  $p(u_{*t})$ . 135

The sand transport rate model proposed by Lettau and Lettau (1978) is adopted because it is widespread in scientific and technical literature (e.g. Fryberger and Dean, 1979; Al-Awadhi and Al-Awadhi, 2009; Barchyn and Hugenholtz, 2011; Yang et al., 2014; Liu et al., 2015), and judged performing better than other sand transport models (Sherman et al., 2013).  $Q_{\theta}$  results from the transformation of the continuous random variables  $u_{*,\theta}$  and  $u_{*t}$ .  $Q_{\theta}$  is expected to be a mixed random variable. In fact,  $Q_{\theta}$  is characterized by a discrete part, i.e.  $Q_{\theta} = 0$ , and a continuous part, i.e.  $Q_{\theta} > 0$ , because of the nature of the adopted sand transport rate model (Eq. 1).

Analytically, given the independent random variables  $u_{*,\theta}$  and  $u_{*t}$ , the probability density function  $f(Q_{\theta})$  for a given value of *d* can be evaluated by differentiating its distribution function  $F(Q_{\theta})$ , which, for  $q \ge 0$  and  $u_{*,\theta} > u_{*t}$ , can be



Figure 2: Threshold shear velocity statistics. Mean values  $\mu(u_{*t})$  and percentiles  $p_1(u_{*t})$ ,  $p_5(u_{*t})$ ,  $p_{25}(u_{*t})$ ,  $p_{75}(u_{*t})$ ,  $p_{99}(u_{*t})$ ,  $p_{99}(u_{*t})$ ,  $p_{99}(u_{*t})$ ,  $p_{10}(u_{*t})$ , p

144 expressed as

$$F_{Q_{\theta}}(q) = P[Q_{\theta} \le q] = P\left[\left\{6.7 \sqrt{\frac{d}{d_{r}}} \frac{\rho_{a}}{g} u_{*,\theta}^{3} \left(1 - \frac{u_{*t}}{u_{*,\theta}}\right) \le q\right\} \bigcap \{u_{*,\theta} > u_{*t}\}\right]$$

$$= \int \int_{\left\{(v_{1}, v_{2}) : v_{1} > v_{2}; v_{1}^{3} \left(1 - \frac{v_{2}}{v_{1}}\right) \le \frac{q}{6.7} \sqrt{\frac{d_{r}}{d}} \frac{g}{\rho_{a}}\right\}} f_{(u_{*,\theta}, u_{*t})}(v_{1}, v_{2}) dv_{1} dv_{2}$$

$$= \int_{0}^{\infty} \left[\int_{v_{1} - \frac{1}{v_{1}^{2}}} \frac{q}{6.7} \sqrt{\frac{d_{r}}{d}} \frac{g}{\rho_{a}}} f_{(u_{*,\theta}, u_{*t})}(v_{1}, v_{2}) dv_{2}\right] dv_{1}$$

$$= \int_{0}^{\infty} \left(F_{u_{*t}}(v_{1}) - F_{u_{*t}}\left(v_{1} - \frac{1}{v_{1}^{2}} \frac{q}{6.7} \sqrt{\frac{d_{r}}{d}} \frac{g}{\rho_{a}}\right)\right) f_{u_{*,\theta}}(v_{1}) dv_{1}.$$
(6)

- However, apart for the untractable analytical solution of this double integration,  $f(Q_{\theta})$  cannot be expressed in explicit
- form because  $f(u_{*t} | d)$  is given by a non-parametric kernel density function (Raffaele et al., 2016).

<sup>147</sup> *Numerically*, Monte Carlo (MC) simulations can be applied (Caflisch, 1998). This approach presents three substantial <sup>148</sup> advantages. First, MC convergence is independent from the number of random variables involved. In fact, it converges <sup>149</sup> with a rate equal to  $1/\sqrt{m}$ , where *m* is the number of realizations, regardless of the number to random variables. <sup>150</sup> Second, the very low cost of each single numerical realization of  $Q_{\theta}$  allows to perform a large number of realizations

for each wind direction. Finally, MC allows to describe the mixed random variable  $Q_{\theta}$  in a straightforward manner.

It is worth pointing out that  $N_{\theta}$  (Eq. 3) is a random quantity because  $T_{\theta}$  is. For this reason, the probability distribution of the directional drift potential  $g(D_{\theta})$  should be expressed as a mixture of convolutions

$$g(D_{\theta}) = \sum_{n=1}^{\infty} (f_1 * \dots * f_i * \dots * f_n) (D_{\theta,\Delta t}) P[N_{\theta} = n] \quad \text{with} \quad f_i = f \quad \text{for} \quad i = 1, \dots, N_{\theta},$$
(7)

whose corresponding mean  $\mu$  and variance  $\sigma^2$  are

$$\mu(D_{\theta}) = \mu(N_{\theta})\mu(D_{\theta,\Delta t})$$

$$\sigma^{2}(D_{\theta}) = \mu(N_{\theta})\sigma^{2}(D_{\theta,\Delta t}) + \mu^{2}(D_{\theta,\Delta t})\sigma^{2}(N_{\theta}).$$
(8)

In particular, the variance results from the sum of two terms, the first due to the variance of  $D_{\theta,\Delta t}$  and the second 155 to the variance of  $N_{\theta}$ . It should be pointed out that adoption of a non-random  $N_{\theta}$  implies an underestimation of 156 the uncertainty of  $D_{\theta}$ , being in this case  $\sigma^2(N_{\theta}) = 0$ . It should be also observed that the constraint on the whole 157 set { $N_{\theta}$ ,  $\theta = 1, 2, ..., n$ } (see Eq. 3) induces a negative dependence between the variables  $N_{\theta}$ , and consequently in 158 the set  $\{D_{\theta}, \theta = 1, 2, \dots, n\}$ , which plays a key role in the final distribution of R. Unfortunately, the description of 159 such an effect of the dependence between the variables  $N_{\theta}$  can not be provided in a simple and tractable analytical 160 manner. For these reasons, a Monte Carlo approach, based on bootstrapping techniques (Efron and Tibshirani, 1993) 161 from the data set of observed values to generate samples, has been adopted. For each simulation, first the vector 162  $N = (N_1, N_2, \dots, N_n)$  of registered occurrences of wind in the considered directions has been obtained from the data 163 set. Then, for each direction  $\{\theta = 1, 2, ..., n\}$ , a sample of cardinality  $N_{\theta}$  of realizations of  $D_{\theta,\Delta t}$  has been randomly 164 chosen. Finally, the matrix  $D = (D_1, D_2, \dots, D_n)$  has been simulated through 165

$$\boldsymbol{D} = \begin{bmatrix} D_1 \\ D_2 \\ \vdots \\ D_{\theta} \\ \vdots \\ D_n \end{bmatrix} = \begin{bmatrix} \sum_{i=1}^{N_1} D_{1,\Delta t,i} \\ \sum_{i=1}^{N_2} D_{2,\Delta t,i} \\ \vdots \\ \sum_{i=1}^{N_{\theta}} D_{\theta,\Delta t,i} \\ \vdots \\ \sum_{i=1}^{N_n} D_{n,\Delta t,i} \end{bmatrix},$$
(9)

where any  $D_{\theta,\Delta t,i}$  is a realization of  $D_{\theta,\Delta t}$  previously extracted from the data set.

<sup>167</sup> Analytically, *R* is the vector sum of the components  $D_{\theta}$  of the matrix *D* (Eq. 4), thus a realization of the resultant <sup>168</sup> drift potential *R* can be immediately assessed once the realization of *D* is given. A set of numerical realizations of *R* <sup>169</sup> can be computed by repeating the same procedure multiple times, and the distribution of *R* can be estimated through <sup>170</sup> such a sample.

#### 171 3. Applications and results

In the following, the proposed SWP approach is applied to five Sites located in the Arabian Peninsula. In Subsection 3.1, the layout of the study is shown. Geographical location and aeolian sand grain size of the chosen sites are reported. In Subsection 3.2, SWP approach is applied to Site 1. Obtained results are shown in terms of both intermediate, i.e.  $Q_{\theta}$  and  $D_{\theta}$ , and final, i.e. *R*, results in order to follow and comment step-by-step the full adopted procedure. Finally, in Subsection 3.3, final results from Sites 1-5 sites are summarized and compared.

#### 177 3.1. Study layout

The site selection obeys to three criteria. Sites with a complete enough anemometric database are first retained. Among them, sites are selected to sample the huge variability of both sand and wind subfields in Arabian Peninsula. Finally, sites are chosen in reason of their proximity to railway lines having in mind the vulnerability of such infras-

<sup>181</sup> tructures to windblown sand.

In Figure 3, Sites 1-5 are represented on Arabian Peninsula (blue dots). On the same Figure, some operating/under construction/planned railway tracks are sketched. In particular, the 950 *km* long Saudi Landbridge links Jeddah with the Saudi Arabia capital Riyadh. The 2750 *km* North South Railway Line links northern Saudi Arabia with Riyadh and the port city Ras Al-Khair. The 450 *km* long Haramain High Speed Rail links the cities of Medina and Mecca. Ethiad Rail is part of the United Arab Emirates' national 1200 *km* railway network.

187 Sites coordinates and mean sand grain size d are reported in Table 1. Mean grain sizes are derived from sedimentology

studies of arabian sand dunes (Al-Sari and Uddin, 1981; Ehlen, 1993; Al-Harthi, 2002; Edgell, 2006). In particular,
 Sites 1, 3 and 4 are sensitive to the sand of Ad Dahna desert, made of medium grained, well sorted quartz sand. Site 2

Sites 1, 3 and 4 are sensitive to the sand of Ad Dahna desert, made of medium grained, well sorted quartz sand. Site 2
 is sensitive to the sand of Jeddah plain. Site 5 is sensitive to the fine grained, moderately well sorted sand of Rub' al

191 Khali desert.

The aerodynamic roughness is set equal to  $z_0 = 4e-3 m$  in all Sites. The wind velocity dataset refers to  $T_r = 5$ years from January 2008 to December 2012 for all stations as well. The 10-min average wind direction is measured in the horizontal plane with a sampling interval  $\Delta \theta = 10^\circ$  at all the selected anemometric stations. n = 36 directions



Figure 3: Sketch of the selected sites (blue dots) and railways tracks (lines)

Site number	Site name	Latitude	Longitude	d [mm]
1	Riyad	24°4'1.20"N	47°34'58.80"E	0.35
2	Jeddah	21°41'60.00"N	39°10'58.80"E	0.25
3	Hafr Al-Batin	27°55'1.20"N	45°31'1.20"E	0.30
4	Al Qaisumah	28°19'1.20"N	46°7'58.80"E	0.30
5	Al Ain	24°12'2.99"N	55°45'40.00"E	0.16

Table 1: Sites of incoming sand drift estimation

result. The 10-min average wind velocity is recorded with a sampling interval in time  $\Delta t = 1$  hour at all the anemometric stations (sampling rate 24/144). The actually available datasets at the selected anemometric stations include missing data due to anemometric breakdowns and/or operational problems. Missing data are in average equal to 4% of the complete dataset. They are evaluated to be almost uniformly distributed along the day. Both the sampling rate

and missing data are sources of incompleteness of the dataset. In literature (see e.g. Burlando et al., 2013) is widely accepted that randomly distributed data incompleteness is usually not influential on the probability distribution of the 10-min average wind velocity, while it may lead to underestimations of the extreme values. It is worth recalling that windblown sand drift potential R is mainly induced by the cumulated values of current values of Q over time, resulting from the 10-min average wind velocity in turn. Hence, data incompleteness is not expected to affect the obtained

results. Finally, the resultant drift potential *R* is expressed over a reference time T = 1 year.

The results discussed in the next Subsections are obtained by MC approach. Hence, results convergence should be discussed every time a random variable is introduced and numerically generated. Convergence is classically evaluated by referring to weighted residuals of the first statistical moments of each random variable. The cardinality of the set of realizations for each random variable is chosen in order to reach a weighted residual lower or at least equal to 1e-2. In the following, the cardinality of each random variable is reported for the sake of completeness, while convergence studies are not reported for the sake of brevity.

#### 211 3.2. Results for site 1

The characteristics of the in-situ sand subfield is summarized by d = 0.35 mm (Table 1). The related input random variable is the threshold shear velocity. Its probability density function  $f(u_{*t}|d = 0.35 mm)$  is derived from Raffaele et al. (2016). Related  $u_{*t}$  statistics are reported in Table 2 in terms of mean value  $\mu$ , standard deviation  $\sigma$  and coefficient of variation *c.o.v.* It is worth recalling that  $f(u_{*t}|d)$  is the same in each wind direction, since it depends solely on sand characteristics.

The wind subfield is obtained by mean wind speed in-situ measurements.  $U_{10}$  variability is assessed in terms of both

<sup>218</sup> *non-directional* and *directional* statistics.

 $_{219}$  Non-directional statistics is summarized in Figure 4.  $U_{10}$  time history is shown in Figure 4(a). Both mean wind speed

 $\mu(U_{10})$  and mean threshold velocity  $\mu(U_t)$  are plotted on the same graph.  $U_{10}$  variability is described by the Hybrid Weibull (HW) model (Takle and Brown, 1978). HW probability density function  $f(U_{10})$  is defined as follows:

$$f_{(\lambda,k)}(U_{10}) = \begin{cases} F_0 & \text{for } U_{10} = 0\\ (1 - F_0) \frac{k}{\lambda} \left(\frac{U_{10}}{\lambda}\right)^{k-1} e^{-U_{10}/\lambda^k} & \text{for } U_{10} > 0 \end{cases}$$
(10)

where  $F_0$  is the rate of zero values, i.e. the frequency of calm wind, k is the shape parameter and  $\lambda$  is the scale parameter. HW  $f(U_{10})$  is plotted in Figure 4(b).



Figure 4: Site 1. Non-directional statistics of mean wind speed: Wind time history (a) and Hybrid Weibull fitting (b)

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Wind shear velocity is recovered from mean wind speed in-situ measurement. HW  $f(U_{10})$  is rescaled into HW  $f(u_*)$ .

 $u_*$  statistical parameters and moments are reported in Table 2, where they can be compared with  $u_{*t}$  ones. In particular,

the threshold shear velocity is higher than the shear velocity in mean terms, while the highest variability is addressed

to the wind subfield.

Table 2: Site 1. Statistical parameters and moments of the non-parametric  $f(u_{*t}|d = 0.35 \text{ mm})$  and Hybrid Weibull  $f(u_{*})$ 

Random variable	$F_0[-]$	k [-]	$\lambda [m \ s^{-1}]$	$\mu [m \ s^{-1}]$	$\sigma [m s^{-1}]$	<i>c.o.v.</i> [–]
$u_{*t}$	-	-	-	0.34	0.06	0.18
$u_*$	0.12	2.09	0.29	0.25	0.13	0.50

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Directional statistics is shown by means of the wind rose and the polar diagram in Figure 5. Calm wind, i.e.  $U_{10,\theta}$ 228 null values, is filtered since it is non-directional by nature and does not contribute in defining directional statistics. 229 Figure 5(a) shows a very broad wind directionality. However, North and South-SouthEast are the directions having 230 the highest occurrence frequency. In Figure 5(b), the empirical probability density function of the wind speed in 231 North direction is shown as an example. Figure 5(c) depicts the variation of probability density function of both wind 232 speed  $U_{10,\theta}$  and erosion threshold  $U_t$  by means of their directional mean values values and extreme percentiles (i.e. 233  $5^{th}$  percentile  $p_5$  and  $95^{th}$  percentile  $p_{95}$ ), as a function of wind direction  $\theta = 1, ..., n$ .  $\mu(U_t)$  is higher than  $\mu(U_{10,\theta})$ 234 for every direction, but the 95<sup>th</sup> percentile of the wind speed  $p_{95}(U_{10,\theta})$  overcomes the corresponding percentile of the 235 threshold velocity  $p_{95}(U_t)$  for winds blowing from around North and from South-SouthWest. 236

 $U_{10,\theta}$  is converted into  $u_{*,\theta}$  dataset. Classic Weibull probability density function  $f(u_{*,\theta})$  are fitted for each direction. Numerical realizations of  $u_{*t}$  and  $u_{*,\theta}$ , consistent with  $f(u_{*t})$  and  $f(u_{*,\theta})$  respectively, are generated in order evaluate the sand transport rate Q within MC approach.  $u_{*t}$  and  $u_{*,\theta}$  cardinality # = 1e+6 is adopted for each direction. Sand transport rate results are organized in the form of sand rose in Figure 6(a) in analogy with the wind rose in Figure 5(a). In fact, the length of each bin is the same in both roses. The wind roses. This is due to the fact that one



Figure 5: Site 1. Directional statistics of mean wind speed: wind rose (a), empirical probability density function of the wind speed in North direction (b), polar diagram of  $U_{10}$  and  $U_t$  statistics (c)

realization of Q for a given direction results from the corresponding realization of  $U_{10}$  along the same direction  $\theta$ . 243 Conversely, the probability density function  $f(Q_{\theta})$  for each direction does not result from a simple rescaling of the 244 corresponding  $f(u_{*,\theta})$ , because of the piece-wise, non-linear transformation (Eq. 1). In particular, for  $0 < u_{*,\theta} < u_{*t}$ , 245  $Q_{\theta} = 0$  even if this does not correspond to wind calm conditions. Hence, the color pattern in each bin significantly 246 varies. An example is explicitly given by the empirical probability density functions for North direction (Fig. 5-b and 247 6-b). Figure 6(c) depicts the mean value and the 95<sup>th</sup> percentile of the sand transport rate as a function of  $\theta$ .  $\mu(Q_{\theta})$  and 248  $p_{95}(Q_{\theta})$  are higher for winds blowing from around North and from South-SouthWest, that are the direction for which 249  $p_{95}(U_{10,\theta}) > p_{95}(U_t)$  (see Fig.5-c).



Figure 6: Site 1. Sand transport rate statistics: sand transport rate rose (a), sand transport rate empirical probability density function in North direction (b), polar diagram of sand transport rate statistics (c)

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The following remarks can be outlined. First, the distribution is no longer a continuous distribution: its hybrid nature 251 is due to the first part of the piece-wise transformation, i.e.  $Q_{\theta} = 0$  if  $u_{*,\theta} \le u_{*t}$ . Second, the distribution is no longer 252 a Weibull-type one, due to the non-linear transformation. In particular, distributions are strongly right-sided skewed. 253

Finally, the sand transport rate directional statistics are strongly bimodal, with North and South prevailing directions, 254 in contrast with the very broad wind directionality (Fig.5-a). This is due to the fact that the sand transport rate  $Q_{\theta}$ 255

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depends on the cube of the effective shear velocity  $u_{*,\theta,eff}^3 = u_{*,\theta}^3 - u_{*,\theta}^2 u_{*t}$ . Referring to Figure 5(c), low-speed winds from West and East do not contribute to  $Q_{\theta}$ , while high-speed winds from North and South almost solely contribute 257

258 to  $Q_{\theta}$ .

The drift potential over the sampling interval  $D_{\theta,\Delta t}$  [ $m^3 m^{-1}hr^{-1}$ ] is simply obtained from  $Q_{\theta}$  [ $Kg m^{-1}s^{-1}$ ] considering the packed bulk sand density  $\rho_b = 1.8e+3 kg m^{-3}$ .

The number of occurrences  $N_{\theta}$  is assessed by bootstrapping a sample of cardinality N = 8768 (i.e. the number of  $\Delta t$ 

in T) from the actual wind velocity dataset. The wind direction frequencies  $N_{\theta}/N$  are shown by box plots in Figure

<sup>263</sup> 7(a). On the same graph, calm wind frequency is plotted too. It should be highlighted that the influence of calm wind

on  $D_{\theta}$  is taken into account by  $N_{\theta}$ . In fact, wind direction frequencies are computed considering the frequency of calm wind (see Eq.3).

Once  $D_{\theta,\Delta t}$  and  $N_{\theta}$  are assessed over each direction, the drift potentials  $D_{\theta}$  over T = 1 year are obtained following

Equation 9. In particular, Equation 9 is applied by bootstrapping a sample of  $D_{\theta,\Delta t}$  and  $N_{\theta}$  realizations, both having

<sup>268</sup> cardinality # = 1e+5. The same cardinality  $\#_{D_{\theta}}$  for each  $D_{\theta}$  follows from MC. In Figure 7(b), drift potential mean

values and percentiles are plotted as a function of  $\theta$  to summarize directional statistics and related  $f(D_{\theta})$ . The non-

parametric probability density function  $f(D_{\theta})$  which describes the incoming sand drift from North in T = 1 year, is shown in Figure 7(c) by way of example.



Figure 7: Site 1. Wind direction frequencies by  $N_{\theta}/N$  box plot (a). Drift potential  $D_{\theta}$  directional statistics (b), drift potential probability density function in North direction (c).

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Three remarks follow. First,  $N_{\theta}$  variability is low, at least for this site. Hence, the variance of  $D_{\theta}$  is mainly due to the variance of  $D_{\theta,\Delta t}$ , while the variance of  $N_{\theta}$  is relatively small (see Eq.8). Second, the drift potential directional statistics are strongly bimodal with North and South prevailing directions in accordance with the sand transport rate ones (see Fig.6-c). Finally, the cumulative sum of the very skewed  $f(Q_{\theta})$  gives rise to almost symmetric  $f(D_{\theta})$ . This is compliant to the central limit theorem: the sum of independent random variables tends to a normally distributed random variable even if the original random variables are not.

Figure 8 provides a synopsis of the uncertainty propagation from erosion threshold and wind speed to sand trans-278 port rate and drift potential. The coefficient of variation and skewness modulus of these random variables are plot-279 ted as a function of the direction  $\theta$  in Figures 8 (a) and (b), respectively. The *c.o.v.* of the input random variables 280  $(U_{10,\theta}, U_t)$  is relatively small (c.o.v.  $\approx 1e-0.5$ ). Uncertainty is magnified by an order of magnitude proceeding to 281  $Q_{\theta}$  (c.o.v.  $\approx 1e+0.5$ ), while c.o.v. is damped again passing from  $Q_{\theta}$  to  $D_{\theta}$  (c.o.v.  $\approx 1e-0.5$ ). Indeed, on the one 282 hand, transformation of random variables done in order to assess Q (i.e. Eq.1) magnifies the uncertainty of the initial 283 random variables  $U_{10,\theta}$  and  $U_t$ . On the other hand, the random sum of identically and independent distributed random 284 variables (Eq.9) damps c.o.v. The c.o.v. of the random variables above shows slight differences over  $\theta$ .  $U_t$  does not 285 depend on  $\theta$  at all, c.o.v.  $(U_{10,\theta})$  is almost constant for this site, c.o.v.  $(Q_{\theta})$  and c.o.v.  $(D_{\theta})$  in turn are higher for winds 286

<sup>287</sup> blowing from East and West, i.e. the less frequent wind directions. The skewness modulus shows approximately the <sup>288</sup> same behavior of *c.o.v.*  $|sk(Q_{\theta})|$  increases significantly with respect to  $|sk(U_t)|$  and  $|sk(U_{10,\theta})|$ , while  $|sk(D_{\theta})|$  decreases <sup>289</sup> again. In particular,  $|sk(D_{\theta})|$  is lower for winds blowing from around North and South directions.



Figure 8: Site 1. Uncertainty propagation from  $U_{10}$  and  $U_t$  to  $Q_\theta$  and  $D_\theta$  in terms of polar diagrams of coefficient of variation (a) and skewness (b)

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Finally, in Figure 9(a), each black dot represents a single realization the resultant drift potential *R*. The radial coordinate of the dots is the vector magnitude |R|, while the angular coordinate is the vector direction  $\hat{R}$ . Each realization of *R* is numerically obtained from one realization of *D* through Equations 4 and 9 by bootstrapping (Efron and Tibshirani, 1993). The ensemble of black dots graphically visualizes the whole set of numerical realizations of *R*. The cardinality of *R* is  $\#_R = 5e+4$ . In the following we call "realization cloud" the ensemble of black dots. The mean resultant drift potential vector  $\mu(R)$  is depicted by the red arrow on the same graph. *R* can be described in probabilistic terms by the joint probability density function  $f(|R|, \hat{R})$  of the two random variables |R| and  $\hat{R}$ . f(|R|) and  $f(\hat{R})$  marginal densities are shown in Figure 9(b) and (c) respectively.

<sup>297</sup> densities are shown in Figure 9(b) and (c), respectively.

The realization cloud appears to be comma-shaped in circular coordinates, i.e. tear-shaped in cartesian coordinates. This shape indicates a significant skewness of  $\hat{R}$ , as testified by its marginal distribution. The radial width of the realization cloud provides a qualitative graphical reading of the variability of *R* magnitude. The circumferential extent of the cloud qualitatively describe the variability of *R* direction. For this site, the variability of  $\hat{R}$  is by far higher than the one of |R|. This is confirmed by the marginal distributions in Figures 9(b) and (c). From a qualitative point of view, it is worth pointing out that the only mean value (red arrow) is a poor description of the sand drift phenomenon. Conversely, realization cloud and related high-order statistics provide a more complete description. In general, SD-

WA approach loses fundamental information of R, while the proposed SWP approach provides complete statistics.

The quantitative statistics of R for this Site and all remaining Sites are reported in the following Subsection.

#### 307 3.3. Comparative analysis Sites 1-5

In the following, all the selected Sites are accounted for. Both wind and windblown sand fields are probabilistically evaluated and critically compared.

In Figure 10,  $U_{10}$  wind roses and polar diagrams of resultant drift potential *R* are represented on Arabian Peninsula map. Realization clouds of the resultant drift potential and marginal densities are plotted as well. On the same graphs, the mean values of *R* are reported (red arrows). In short, Figure 10 collects the results of the initial and final step of the proposed procedure. Wind rose shape testifies a variety of wind regimes: wide unimodal, i.e. Site 2, acute bimodal, i.e. Site 3, obtuse bimodal, i.e. Sites 1 and 5, complex, Site 4. Realization cloud shape, dimension and density change significantly moving from a site to another. The realization clouds appear to be comma-shaped in polar diagrams (i.e. tear-shaped in cartesian coordinates), or kidney-shaped (i.e. elliptical-shaped in cartesian coordinates). Comma



Figure 9: Site 1. Resultant drift potential (a), resultant drift potential magnitude marginal density (b) and resultant drift potential direction marginal density (c).

shape (Sites 1 and 5) indicates a significant skewness of  $\hat{R}$ , while kidney shape (Sites 2,3,4) indicates weakly skewed 317 magnitude and direction. The wider the realization cloud in radial and/or circumferential direction, the higher the 318 variation of R, in magnitude and direction, respectively. Variations of both |R| and  $\hat{R}$  are small at Sites 2 and 3, so 319 that kidney-shaped cloud appear as elliptical. Site 4 is remarkably characterized by very high variation of  $\hat{R}$  and a 320 small variation of |R|. The marginal densities f(|R|) and  $f(\hat{R})$  clearly reflect these differences. In particular, while in 321 some cases they recall Gaussian distributions (Sites 2 and 3), in others they appear asymmetric, mainly with respect 322 to the direction ( $f(\hat{R})$ ) at Sites 1 and 5). In general, the relation between wind rose and realization cloud is not straight-323 forward, because of the non-linear relation between  $U_{10}$  and Q. Furthermore, wind roses graphically point out wind 324 direction frequencies much more effectively than wind speed frequencies. However, it is worth pointing out that the 325 more complex the wind rose, the wider the realization cloud. 326

Non-dimensional statistics of both |R| and  $\hat{R}$  are reported in Table 3 to summarize the obtained results and quan-327 titatively compare the Sites. Variation and skewness of |R| and  $\hat{R}$  are assessed in order to understand how much the 328 random variables are dispersed and how far are from Gaussianity. The variability of |R| is expressed by means of 329 *c.o.v.*, while the variability of  $\hat{R}$  is directly expressed by the angular deviation  $\sigma$ . It is worth to point out that since  $\hat{R}$ 330 is a circular random variable, circular statistics is assessed (Fisher, 1995; Berens, 2009). The lowest variability is ad-331 dressed to Site 2, i.e. the Site with unimodal wind regime, while the highest variability is addressed to e.g. Sites 1 and 332 4, i.e. the Sites with obtuse bimodal or complex wind regimes. Concerning probability density functions symmetry, 333 Sites 1 and 5 show the most skewed distributions, while Site 3 one is almost symmetric. 334

The design of infrastructures in arid environments should be based on sand drift magnitude related to a low prob-335 ability of exceedance. Hence, characteristic values (i.e. extreme percentiles) of both R magnitude and direction are 336 included in Table 3. The ratio between 95<sup>th</sup> percentile and mean value  $p_{95}/\mu$  is assessed as regards R magnitude. The 337 study gives rise to characteristic values up to  $\approx 1.6$  times the mean value (Site 1). In other words, the evaluation of 338 |R| in mean terms only significantly underestimates the amount of transported sand. The angular distance  $|p_{95} - p_5|$  is 339 evaluated, regarding  $\hat{R}$ . Both percentiles are referred to anti-clockwise circular direction from East. In other words, 340  $|\mathbf{p}_{05} - \mathbf{p}_5|$  provides a quantitative measure of the variability of  $\hat{R}$  based on characteristics directions. This measure is the 341 well posed probabilistic reading of the estimate of drift direction variability proposed by Fryberger and Dean (1979) 342



Figure 10: Sites 1-5. Wind roses and resultant drift potentials around Arabian Peninsula 13

in deterministic terms through Equation 5. The highest  $|p_{95} - p_5|$  angular distance is observed for Site 4, i.e.  $\approx 88^\circ$ , while the lowest,  $\approx 8^\circ$ , is observed for Site 2.

	$ R  [m^3 m^{-1} y r^{-1}]$			$\hat{R}$ [°]			
	c.o.v. [-]	sk [-]	$p_{95}/\mu[-]$	-	σ[°]	sk [-]	$ p_{95} - p_5 [^\circ]$
Site 1	0.32	0.37	1.57		20.77	2.61	77.22
Site 2	0.06	0.02	1.09		2.52	0.38	8.28
Site 3	0.10	0.01	1.16		3.88	-0.05	12.78
Site 4	0.27	0.11	1.46		24.09	0.60	87.80
Site 5	0.10	0.28	1.17		7.11	1.05	23.22

Table 3: Sites 1-5. Statistics of resultant drift potential magnitude and direction

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#### 345 4. Conclusions

The present study introduces a new Sand-Wind Probabilistic (SWP) approach to evaluate incoming windblown sand drift potentials and resultant drift potentials. The approach adapts the general framework proposed by Fryberger and Dean (1979) in order to deal with the sources of uncertainty related to both wind and sand subfields. The input uncertainties on  $U_{10}$  and  $u_{*t}$  propagate to the final result, i.e. *R*, passing through the definition of  $Q_{\theta}$  and  $D_{\theta}$ .

The following concluding remarks can be outlined, bearing in mind the three kickoff questions raised in Intro-350 duction. First, uncertainty of both threshold shear velocity and mean wind velocity are magnified passing to the 351 directional sand transport rate  $Q_{\theta}$  by about an order of magnitude. Subsequently, uncertainty is damped from  $Q_{\theta}$  to 352 the drift potential  $D_{\theta}$ , and it is further damped to the resultant drift R. Magnification is due to the cubic dependency 353 of Q versus  $u_*$  and  $u_{*t}$ , while damping results from cumulating in time and vector summing over directions. Second, 354 the probability distribution of the resultant drift potential changes significantly form a site to another in the same 355 region. Complex wind regimes are particularly prone to cause windblown sand drift with high inborn variability. For 356 instance, the highest c.o.v(|R|) and  $\sigma(R)$  are referred to sites showing obtuse bimodal or complex wind roses. Changes 357 in the sand granulometry and related shear threshold velocity probability distribution from one site to another also 358 affect R. Finally, the proposed SWP probabilistic approach allows to obtain characteristic values of R, while the Sand 359 Deterministic-Wind Averaged (SD-WA) approach adopted up to now in scientific literature and engineering practice 360 does not provide sufficient statistics to describe correctly the phenomenon. The gap between characteristic and mean 361 value of RDP makes the approach of interest for engineering practice and grounds the semi-probabilistic approach 362 to design of civil infrastructure in arid regions. Regarding sites with complex wind regimes, on the one hand, the 363 characteristic value of |R| is about 1.5 times the mean value. On the other hand, the angular distance between the mean 36 direction and the characteristic values of  $\hat{R}$  is about 40°. 365

In the light of the obtained results, we suggest four research perspectives. First, the role played by each considered 366 random variable in variability of sand drift should be ascertained by means of a numerical sensitivity study, i.e. by 367 setting a constant grain diameter (and hence the probability density function of  $u_{*t}$ ) and varying the wind field, and 368 vice versa. Second, the proposed approach needs to be validated by in-situ, long-term, continuous and automatic 369 recording of the sand drift, analogously to wind speed measurements. Traditional sand trap (Nickling and Neuman, 370 1997; Weaver and Wiggs, 2011) are not adequate to this purpose. Piezoelectric sand flux sensor (Udo, 2009) proved 371 encouraging performances during prototype testing in operational conditions. This technology will enable in the next 372 future years-long site measurements. Third, having in mind the vast amount of sand transport rate models in litera-373 ture, Q model uncertainty should be investigated and, if significant, incorporated in the adopted probabilistic method. 374 Finally, sand drift extreme values statistics would worth to be described in order to assess how much stand storms 375 weigh on the total amount of the resultant drift potential and on disastrous events. 376

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